
Song Xi Chen
Iowa State University, songchen@iastate.edu

Hengjian Cui
Beijing Normal University

Follow this and additional works at: http://lib.dr.iastate.edu/stat_las_preprints
Part of the Statistics and Probability Commons

Recommended Citation
http://lib.dr.iastate.edu/stat_las_preprints/36

Song Xi Chen and Hengjian Cui

1 Department of Statistics, Iowa State University and

2 Department of Mathematics, Beijing Normal University

Abstract. This paper considers the second order properties of empirical likelihood for a parameter defined by moment restrictions, which is the framework operated upon by the Generalized Method of Moments. It is shown that the empirical likelihood defined for this general framework still admits the delicate second order property of Bartlett correction, which represents a substantial extension of all the established cases of Bartlett correction for the empirical likelihood. An empirical Bartlett correction is proposed, which is shown to work effectively in improving the coverage accuracy of confidence regions for the parameter.

Key words: Bartlett correction, Coverage accuracy, Empirical likelihood, Generalized Method of Moments.

1. Introduction

Generalized Method of Moments (GMM) introduced by Hansen (1982) is an important inferential framework in econometric studies. GMM is based on, upon given a model, some known functions $g(X, \theta)$ of a random observation $X \in \mathbb{R}^d$ and an unknown parameter $\theta \in \mathbb{R}^p$, where $g : \mathbb{R}^{d+p} \rightarrow \mathbb{R}^r$, such that $E\{g(X, \theta)\} = 0$ which constitutes moment restrictions on the relationship between $X$ and $\theta$. The power of GMM is in its allowing $r \geq p$, namely the

---

1The research of this paper was supported by a National University of Singapore Academic Research Grant (R-155-000-018-112) and a RFDP of China grant (20020027010). The authors thank Dr Weidong Lin for computational support, who was supported under the National University of Singapore Academic Research Grant.
number of moment restrictions (instruments) can be larger than the number of parameter, which leads to a full exploration of inference opportunities provided by the given model. There is a vast pool of literatures on GMM. Here we only cite the latest reviews of Andrews (2002), Brown and Newey (2002), Imbens (2002) and Hansen and West (2002).

Empirical likelihood (EL) introduced by Owen (1988) is a computer-intensive statistical method that facilitates a likelihood-type inference in a nonparametric or semiparametric setting. It is closely connected to the bootstrap as the EL effectively carries out the resampling implicitly. On certain aspects of inference, EL is more attractive than the bootstrap, for instance its ability of internal studentizing so as to avoid explicit variance estimation and producing confidence regions with natural shape and orientation; see Owen (2001) for an overview of EL. A key property of EL is that the log EL ratio is asymptotically chi-squared distributed, which resembles the Wilks’ theorem in parametric likelihood. The Wilks’ theorem was established in the original proposal of Owen (1988) for the means, in Hall and La Scala (1990) for smoothed function of means, Qin and Lawless (1994) for parameters defined by moment restrictions and Kitamura (1997) for weakly dependence observations.

There have been comprehensive studies of EL in the context of GMM in econometrics. Imbens (1997) shows that the maximum EL estimator of $\theta$ is a one-step variation of the two-stage GMM estimator in the over-identified case of $r > p$, and achieves the same asymptotic efficiency as the two-stage estimator. Testing is considered in Kitamura (2001) for moments restrictions, and Tripathi and Kitamura (2002) for conditional moment restrictions. Estimation and testing with conditional moment restrictions are studied in Donald, Imbens and Newey (2003) and Kitamura, Tripathi and Ahn (2002). They found that EL posses the attractive features of avoiding estimating optimal instruments and achieving asymptotic pivotalness. Tilted EL and other variations are studied in Kitamura and Stutzer (1997), Smith (1997) and Newey and Smith (2004). In particular, Newey and Smith (2004) find that the EL estimator is favorable in terms of the bias and the second order variance in comparison with the GMM estimator.
Another key property of the EL is Bartlett correction, which is a delicate second order property implying that a simple mean adjustment to the likelihood ratio can improve the approximation to the limiting chi-square distribution by one order of magnitude and hence can be used to enhance the coverage accuracy of likelihood-based confidence regions. In the context of testing hypotheses, the Bartlett correction reduces the errors between the nominal and actual significant levels of an EL test. Bartlett correction has been established for EL by DiCiccio, Hall and Romano (1991) for smoothed functions of means and Chen (1993, 1994) for linear regression. Baggerly (1998) shows that EL is the only member within the Cressie-Read power divergence family that is Bartlett correctable. Jing and Wood (1996) reveal that the exponentially tilted EL for the means is not Bartlett correction as the tilting alters the delicate second order mechanism of EL.

In this paper we show that the EL with moment restrictions is still Bartlett correctable. The finding represents a substantial extension to the established cases of Bartlett correction, which almost all of them assume \( r = p \) corresponding to the just-identified case in GMM. The establishment of the Bartlett correction for the just-identified case is a lot easier as the log maximum EL takes a constant value \(-n \log(n)\) (\(n\) is the sample size). However, in the over-identified case the maximum EL is no longer a constant, rather it introduces many extra terms into the log EL ratio and makes the study of Bartlett correction far more challenging as can be seen from the analysis carried out in this paper. The establishment of Bartlett correction in this general case indicates that EL inherits the delicate second order mechanism of the parametric likelihood in a much wider situation. This together with the findings of Imbens (1997), Kitamura (2001) and Newey and Smith (2004) and others suggests that the EL is an attractive inferential tool in the context of moment restrictions. The establishment of the Bartlett correction leads to a practical Bartlett correction, which is confirmed to work effectively for coverage restoration in our simulation studies reported in Section 4.

The paper is organized as follows. Section 2 provides an expansion for the log EL ratio for parameters defined by moment restrictions. Bartlett correction and coverage errors
assessment of EL confidence regions are investigated in Section 3. Simulation results are reported in Section 4, followed by a general discussion in Section 5. All technical details are left in the appendix.

2. EL for generalized moment restrictions

Let $X_1, X_2, \ldots, X_n$ be $d$-dimensional independent and identically distributed random sample whose distribution depends on a $p$-dimensional parameter $\theta$ which takes values in a compact parameter space $\Theta \subseteq \mathbb{R}^p$. The information about $\theta$ is summarized in the form of $r$ unbiased moment restrictions $g^j(x, \theta)$, $j = 1, 2, \ldots, r$, such that $E[g^j(X_1, \theta_0)] = 0$ for a unique $\theta_0$, which is the true value of $\theta$. Let

$$g(X, \theta) = (g^1(X, \theta), g^2(X, \theta), \ldots, g^r(X, \theta))^T \quad \text{and} \quad V = Var\{g(X_1, \theta_0)\}.$$

We assume the following regularity conditions:

(2.1) (i) $V$ is a $r \times r$ positive definite matrix and the rank of $E[\partial g(X_1, \theta_0)/\partial \theta]$ is $p$;

(ii) For any $j$, $1 \leq j \leq p$, all the partial derivatives of $g^j(x, \theta)$ up to the third order with respect to $\theta$ are continuous in a neighborhood of $\theta_0$ and are bounded by some integrable functions respectively in the neighborhood;

(iii) $\limsup_{|t| \to \infty} |E[\exp\{it^T g(X_1, \theta_0)\}]| < 1$ and $E\|g(X_1, \theta_0)\|^{15} < \infty$.

Conditions (i) and (ii) are standard requirements for establishing the Wilks’ theorem and higher order Taylor expansions of the EL ratio. The first part of the condition (iii) is just the Cramér’s condition on the characteristic function of $g(X, \theta_0)$. It and the requirement that $E\|g(X, \theta_0)\|^{15} < \infty$ are required for establishing the Edgeworth expansion.

To facilitate simpler expressions, we transform $g(X_i, \theta)$ to $w_i(\theta) = TV^{-1/2}g(X_i, \theta)$ where $T$ is a $r \times r$ orthogonal matrix such that

$$(2.2) \quad TV^{-1/2}E\left(\frac{\partial g(X_i, \theta_0)}{\partial \theta}\right)U = \left(\Lambda, 0\right)^T_{r \times p}.$$

Here $U = (u^{kl})_{p \times p}$ is an orthogonal matrix and $\Lambda = diag(\lambda_1, \cdots, \lambda_p)$ is non-singular.
Let \( p_1, p_2, \cdots, p_n \) be non-negative weights allocated to the observations. The EL for \( \theta \) as proposed in Qin and Lawless (1994) is \( L(\theta) = \prod_{i=1}^{n} p_i \) subject to \( \sum_{i=1}^{n} p_i = 1 \) and \( \sum_{i=1}^{n} p_i w_i(\theta) = 0 \). Let \( \ell(\theta) = -2\log \{L(\theta)/n^n\} \). Standard derivations in EL show

\[
\ell(\theta) = 2 \sum_{i=1}^{n} \log \{1 + \lambda^T w_i(\theta)\}
\]

where \( \lambda = \lambda(\theta) \) is the solution of \( n^{-1} \sum_{i=1}^{n} \frac{w_i(\theta)}{1 + \lambda^T w_i(\theta)} = 0 \). According to Qin and Lawless (1994), the maximum EL estimator \( \hat{\theta} \) and its corresponding \( \lambda \), denoted as \( \lambda^* \), are solutions of

\[
Q_{1n}(\lambda, \theta) = n^{-1} \sum_{i=1}^{n} \frac{w_i(\theta)}{1 + \lambda^T w_i(\theta)} = 0 \quad \text{and}
\]

\[
Q_{2n}(\lambda, \theta) = n^{-1} \sum_{i=1}^{n} \frac{(\partial w_i(\theta)/\partial \theta)^T \lambda}{1 + \lambda^T w_i(\theta)} = 0.
\]

Then the log EL ratio is \( r(\theta) = \ell(\theta) - \ell(\hat{\theta}) \).

In the following we are to develop expansions to \( \ell(\theta_0) \) and \( \ell(\hat{\theta}) \) respectively. To expand \( \ell(\theta_0) \), define

\[
\alpha^{j_1 \cdots j_k} = E\{w_{i_1}^{j_1}(\theta_0) \cdots w_{i_k}^{j_k}(\theta_0)\} \quad \text{and}
\]

\[
A^{j_1 \cdots j_k} = n^{-1} \sum_{i=1}^{n} w_{i_1}^{j_1}(\theta_0) \cdots w_{i_k}^{j_k}(\theta_0) - \alpha^{j_1 \cdots j_k}.
\]

Here we use \( a^j \) to denote the \( j \)-th component of a vector \( a \). Then, it may be shown that

\[
n^{-1}\ell(\theta_0) = A^j A^j - A^{j_1} A^{j_1} A^i + \frac{2}{3} \alpha^{i j_1} A^j A^i A^h + A^{i j} A^{i j} + A^{j_1} A^{i j} A^h + \frac{2}{3} A^{i j} A^j A^i A^h - 2\alpha^{i j_1} A^j A^i A^h A^g + O_p(n^{-5/2}).
\]

We use here a convention where if a superscript is repeated a summation over that superscript is understood. This expansion has the same form as DiCiccio, Hall and Romano (1991) for the mean parameter when \( r = p \) and Chen (1993) for linear regression.

It is quite challenging to expand \( \ell(\hat{\theta}) \) in the general case of \( r > p \). Two new systems of notations are introduced to facilitate the expansion. Let \( \eta = (\lambda, \theta), Q(\eta) = (Q_{1n}(\eta), Q_{2n}(\eta))^T \), \( S_{21} = U(\Lambda, 0) \) and \( S_{12} = S_{21} \). Due to the early transformation in (2.2),

\[
S =: E\{\frac{\partial Q(0, \theta_0)}{\partial \eta}\} = \begin{pmatrix} -I & S_{12} \\ S_{21} & 0 \end{pmatrix}.
\]
Put $\Gamma(\eta) = S^{-1}Q(\eta)$. Now we can introduce the notations involving $\Gamma(\eta)$ and their derivatives

$$\beta_{j_1j_2\cdots j_k} = E\left( \frac{\partial^k \Gamma^j(0, \theta_0)}{\partial \eta_{j_1}\cdots \partial \eta_{j_k}} \right) \quad \text{and} \quad B_{j_1j_2\cdots j_k} = \frac{1}{n} \sum_{i=1}^{n} \frac{\partial^k \Gamma^j(0, \theta_0)}{\partial \eta_{j_1i}\cdots \partial \eta_{j_ki}} - \beta_{j_1j_2\cdots j_k}$$

and the notations involving $w_i(\theta)$ and their derivatives

$$\gamma_{j_1j_2\cdots j_k;k_1k_2\cdots k_m;p_1p_2\cdots p_n} = E\left( \frac{\partial^m w_i(\theta_0)}{\partial \theta_{j_1}\partial \theta_{j_2}\cdots \partial \theta_{j_k}} \frac{\partial^p w_i(\theta_0)}{\partial \theta_{k_1}\cdots \partial \theta_{k_m}} \frac{\partial^p w_i(\theta_0)}{\partial \theta_{p_1}\cdots \partial \theta_{p_n}} \right) \quad \text{and} \quad C_{j_1j_2\cdots j_k;k_1k_2\cdots k_m;p_1p_2\cdots p_n} = \frac{1}{n} \sum_{i=1}^{n} \gamma_{j_1j_2\cdots j_k;k_1k_2\cdots k_m;p_1p_2\cdots p_n} - \gamma_{j_1j_2\cdots j_k;k_1k_2\cdots k_m;p_1p_2\cdots p_n}.$$
(2.8) + 2\alpha^{ijh} B^i B^i B^h B^q - \alpha^{ijh} \beta^{i,jq} B^i B^j B^q B^h - \frac{1}{2} \alpha^{jihg} B^j B^j B^h B^g + O_p(n^{-5/2})

where [2, j, i] indicates there are two terms by exchanging the super-scripts i and j, and the same is understood for [3, i, j, k]. Expansion (2.8) for \( \ell(\hat{\theta}) \) is much more complicated than the just-identified case of \( r = p \). In that case, all the \( B^j = 0 \) from a result established in (A.1), which means \( \ell(\hat{\theta}) = 0 \) and \( r(\theta_0) = \ell(\theta_0) \). This is the situations of almost all the existing studies on Bartlett corretability of the EL. When \( r > p \), the expansion of \( \ell(\hat{\theta}) \) contains more terms than that of \( \ell(\theta_0) \), which increases substantially the difficulty of the second order study.

Combining (2.5) and (2.8), and carrying out further simplifications,

\[
n^{-1}r(\theta_0) = A^i A^i - A^{kl} A^k A^l - 2A^i p+a A^{p+a} A^l + \frac{2}{3} \alpha^{klm} A^k A^m A^l + 2\omega^{kl} C^{p+a,k} A^{p+a} A^l + (2\alpha^{kl} p+a - \gamma^{p+a,mm} w_{mk} w_{nl}) A^{p+a} A^k A^l + A^{ji} (A^{kl} A^l A^h - B^{i,q} B^q B^j[2, i, j]) + 2 \left( \alpha^{p+aq+b} - \gamma^{p+a,p+b,k} \omega^{kl} \right) A^{p+a} A^{p+b} A^l + B^{i,a} B^{j,a} B^s B^q + 2\gamma^{j,i} B^i B^i B^r B^s B^q - 2\gamma^{j,i} B^i B^i B^r B^s B^q - 2\alpha^{jih} B^j B^j B^h B^q + 2\gamma^{j,i} (B^i B^i B^r B^q + B^i B^i B^r B^q B^q[2, j, i]) + (\gamma^{j,i,ik} + \gamma^{j,i,l,ik}) B^j B^i B^r B^s B^q + (\frac{1}{2} \alpha^{i,jm} \gamma^{j,sl} - \frac{1}{2} \gamma^{i,jm} \beta^{j,s,k} \delta^{j,k} \mu^{j,sl} \gamma^{j,k}) B^i B^s B^i B^s B^q + (\alpha^{ijh} \beta^{i,jh,uk} - \gamma^{j,i,ik,il} \beta^{j,s,l} \gamma^{j,k}) B^i B^j B^h B^q + (\gamma^{j,k} \beta^{j,r,l} \gamma^{j,uk} - \gamma^{j,i,jk} \beta^{j,uk} \gamma^{j,k}) B^i B^h B^i B^h B^q - \frac{1}{2} \gamma^{j,kl} \beta^{j,kl} \mu^{j,uk} \gamma^{j,k} B^i B^k B^s B^r B^s B^q + C^{j,k} B^i B^i B^r B^s B^q - 2C^{j,i} B^i B^i B^r B^s + \frac{2}{3} A^{jih} (B^j B^h B^h + A^i A^i A^h B^q + O_p(n^{-5/2})
\]

This expansion leads to the following signed root decomposition

\[
n^{-1}r(\theta_0) = (R_1 + R_2 + R_3)^j (R_1 + R_2 + R_3)^j + O_p(n^{-5/2})
\]

7
where $R_i = O_p(n^{-i/2})$ for $i = 1, 2$ and 3. Clearly, the terms appeared in the first two lines of (2.9) fully determine $R_1$ and $R_2$, namely

$$R_1^l = A^l,$$

$$R_2^l = -\frac{1}{2} A^{kl} A^k - A^l p + a A^{p+a} + \frac{1}{3} \alpha klm A^k A^m + \omega^{kl} C^{p+a,k} A^{p+a} + \left[ \alpha^{kl} p^{+a} + \frac{1}{2} \gamma^{p+a,mm} \omega^{kl} \right] A^{p+a} A^k + \left[ \alpha^l p^{+a,p+b} - \gamma^{p+a,p+b,k} \omega^{kl} \right] A^{p+a} A^{p+b}$$

and $R_1^l = R_2^l = 0$ for $j \in \{p + 1, \ldots, r\}$. An expression for $R_3^l$ is given in Appendix 2.

From (2.9), $r(\theta_0) = nA^l A^l + o_p(1)$ which means that $r(\theta_0) \xrightarrow{d} \chi^2_p$ and leads to an EL confidence region for $\theta$ with nominal confidence level $1 - \alpha$: $I_\alpha = \{\theta | r(\theta) \leq c_\alpha\}$ where $c_\alpha$ is the upper $\alpha$-quantile of $\chi^2_p$ distribution.

3. The second order properties

The coverage accuracy of the EL confidence region $I_\alpha$ is evaluated in the following theorem.

**Theorem 1.** Under conditions (2.1),

$$P\{r(\theta_0) < c_\alpha\} = \alpha - n^{-1} p^{-1} B_c c_\alpha f_p(c_\alpha) + O(n^{-2})$$

where $f_p(\cdot)$ is the density of $\chi^2_p$ distribution, $B_c = p^{-1} (\sum_{l=1}^p \Delta^l + \frac{1}{36} \alpha^{lkk} \alpha^{lmm})$ and $\Delta^l$ is defined in (A.18).

Theorem 1 indicates that the coverage error of the EL confidence region $I_\alpha$ is $O(n^{-1})$, which is the same order as a standard two sided confidence region based on the asymptotic normality of $\hat{\theta}$. The attractions of the EL confidence region are (i) there is no need to carry out any secondary estimation procedure in formulating the confidence region whereas the covariance matrix has to be estimated for the confidence region based on the asymptotic normality; and (ii) the shape and the orientation of the region are naturally determined by the likelihood surface, free of any subjective intervention.
Theorem 2. Under conditions (2.1),

\[ P\{r(\theta_0) < c_n(1 + n^{-1}B_c)\} = \alpha + O(n^{-2}). \]

The theorem shows that the Bartlett correction is maintained by the EL for the situation of general moment restrictions, despite that \( r \) may be larger than \( p \) and \( \ell(\hat{\theta}) \) has a rather complex expression. This indicates that the EL is resilient in sharing this delicate second order property with a parametric likelihood and the existence of certain internal mechanism in the EL that resembles that of the parametric likelihood.

It can be seen from Appendix A.4. that the Bartlett factor \( B_c \) has a rather involved expression for a general over-identified case of \( r > p \) due to the lengthy expressions of \( \Delta ll \). However, it admits simpler expression in two special situations. One is in the situation of just-identified moment restrictions with \( r = p \). It may be easily checked from (A.18) that

\[ B_c = p^{-1} \left( \frac{1}{2} \alpha^{llk} - \frac{1}{3} \alpha^{lkm} \alpha^{km} \right), \]

which is the Bartlett factor obtained in DiCiccio, Hall and Romano (1991) for smooth function of means and Chen (1993) for linear regression. The other situation is when \( r > p \), but (i) \( Cov\{g^l(X, \theta_0), g^{p+a}(X, \theta_0)\} = 0 \) for any \( j \leq p \) and \( a \leq r - p \) and (ii) \( g^l(x, \theta) = g^l(x) \) does not depend on \( \theta \) for \( j = p + 1, \cdots, r \). The assumption (i) means that the first \( p \) estimating equations are uncorrelated with the last \( r - p \) estimating equations at \( \theta_0 \) and (ii) means that the last \( r - p \) estimating equations are free of parameters. In this case,

\[ B_c = p^{-1} \left( \frac{1}{2} \alpha^{llk} + \alpha^{ll} p+a p+a - \frac{1}{3} \alpha^{lkm} \alpha^{km} - \alpha^{llk} \alpha^{k+p+a} p+a - \alpha^{lf} p+a \alpha^{lf} p+a \right). \]

To practically implement the Bartlett correction in a general situation, \( \hat{B}_c =: 1 + B_c n^{-1} \) has to be estimated. It is noted that the direct plug-in estimator of \( \hat{B}_c \) can be obtained by substituting all the populations moments involved by their corresponding sample moments. However, considering the rather lengthy forms of \( \hat{B}_c \), we propose using the bootstrap to estimate \( \hat{B}_c \), which is based on the fact that

\[ E\{r(\theta_0)\} = p(1 + B_c n^{-1}) + O(n^{-2}) = p\hat{B}_c + O(n^{-2}) \]

(3.1)
by combining the expressions of $E(R_i^lR_j^r)$ given in Appendix 3. The bootstrap procedure is

Step 1: generate a bootstrap resample $\{X_i^*\}_{i=1}^n$ by sampling with replacement from the original sample $\{X_i\}_{i=1}^n$ and compute $r^*(\hat{\theta}) = \ell^*(\hat{\theta}) - \ell^*(\hat{\theta}^*)$, where $\ell^*$ and $\hat{\theta}^*$ are respectively the Log EL ratio and the maximum EL estimate based on the resample;

Step 2: for a large integer $N$, repeat Step 1 $N$ times and obtain $r_1^*(\hat{\theta}), \ldots, r_N^*(\hat{\theta})$.

As $N^{-1} \sum_{b=1}^N r^*_b(\hat{\theta})$ estimates $E\{r(\theta_0)\}$, a bootstrap estimate of $\hat{B}_c$ is

$$\hat{\hat{B}}_c = (Np)^{-1} \sum_{b=1}^N r^*_b(\hat{\theta}).$$

Let $X_n = \{X_1, \ldots, X_n\}$ be the original sample. It may be shown by standard bootstrap arguments, for instance those given in Hall (1992), that

$$E\{\hat{\hat{B}}_c|X_n\} = (1 + B_cn^{-1})\{1 + O_p(n^{-1/2})\}$$

which means that the bootstrap estimate of $\hat{\hat{B}}_c$ is $\sqrt{n}$-consistent. Now a practical Bartlett corrected confidence region is $I_{\alpha,bc} = \{\theta| r(\theta) \leq c_\alpha \hat{\hat{B}}_c\}$. It can be shown from Theorem 2 and (3.2) that the coverage error of $I_{\alpha,bc}$ is $O(n^{-3/2})$, which improves that of $I_\alpha$.

The above use of the bootstrap to estimate the Bartlett factor $B_c$ or $\hat{\hat{B}}_c$ naturally leads ones to think of using the bootstrap to calibrate directly on the distribution of the EL ratio $r(\theta_0)$. Let us rank $\{r^i(\hat{\theta})\}_{i=1}^N$ such that $r^1(\hat{\theta}) \leq r^2(\hat{\theta}) \leq \cdots \leq r^N(\hat{\theta})$. Then a direct bootstrap confidence interval with a nominal level $1 - \alpha$ is $I_{\alpha,bl} = (r^*([\alpha N/2]+1)(\hat{\theta}), r^*([1-\alpha N/2]+1)(\hat{\theta}))$ where $[\cdot]$ is the integer truncation operator.

The cumulants and expansions which are quite expensively derived for the purpose of establishing Bartlett correction are needed in assessing the coverage accuracy of the direct Bootstrap confidence interval $I_{\alpha,bl}$. It may be shown that under conditions (2.1),

$$P(\theta_0 \in I_{\alpha,bl}) = \alpha + O(n^{-3/2})$$

We would not provide the proof here due to a limited space. It can be carried out by taking a similar route as in Hall (1992) and utilizing the Edgeworth expansion established in the proof of Theorem 1.
which indicates that the coverage error of the bootstrap confidence interval $I_{\alpha, bt}$ and the Bartlett corrected interval $I_{\alpha, bc}$ is at the same order. This is indeed confirmed by our simulation studies reported in the next section, although we observe that the performance of the Bartlett corrected interval is more robust.

4. Simulation Results

We report in this section results of two simulation studies designed to confirm the theoretical finding of Bartlett correction of the EL by implementing the proposed empirical Bartlett correction. For comparison purposes, the bootstrap confidence intervals $I_{\alpha, bt}$ is also evaluated.

In the first simulation study, $X_1, \cdots, X_n$ are independent and identically $N(\theta, \theta^2 + 1)$ distributed, as considered in an example of Qin and Lawless (1994). The relationship between the mean and variance leads to moment restrictions: $g_1(X_1, \theta) = X_1 - \theta$ and $g_2(X_1, \theta) = X_1^2 - 2\theta - 1$. This is an over-identified case as there are two moment restrictions and one parameter of interest, i.e. $r = 2$ and $p = 1$. Like Qin and Lawless, the value of $\theta$ is chosen to be 0 and 1 respectively. The sample size used in the simulation study is $n = 20, 30, 40$ and 50 respectively.

In the second simulation study, we consider the following autoregressive panel data model, which is an example considered in Brown and Newey (2002)

$$X_{it} = \theta X_{it-1} + \alpha_i + \epsilon_{it}, \quad X_{i0} = \frac{\alpha_i}{1 - \rho} + \epsilon_i,$$

for $t = 1, \cdots, 4$ and $i = 1, \cdots, n$, where $|\theta| < 1$, $\{\epsilon_{it}\}_{t=1}^4$ and $\alpha_i$ are mutually independent standard normal random variables, $v_i \sim N(0, (1 - \theta^2)^{-1})$ and independent of $\{\epsilon_{it}\}_{t=1}^4$ and $\alpha_i$. Let $X_i = (X_{i1}, \ldots, X_{i4})$. The moment restrictions after taking time differencing are $g_1(X_i, \theta) = X_{i1}(\Delta X_{i3} - \theta \Delta X_{i2}), g_2(X_i, \theta) = X_{i1}(\Delta X_{i4} - \theta \Delta X_{i3})$ and $g_3(X_i, \theta) = X_{i2}(\Delta X_{i3} - \theta \Delta X_{i2})$ where $\Delta X_{it} = X_{it} - X_{it-1}$. It is easy to check from model (4.1) that $E\{g_j(X_i, \theta)\} = 0$. Hence, there are three constraints and one parameter, i.e. $r = 3$ and $p = 1$, another over-identified case. The parameter $\theta$, which is the autoregressive coefficient is assigned values
of 0.5 and 0.9 to obtain different levels of correlations. The sample size is chosen at \( n = 50 \) and 100 respectively.

In both simulation studies, the empirical coverage and length of the EL, Bartlett corrected EL and the direct bootstrap calibrated intervals are evaluated with nominal coverage levels of 90% and 95% respectively. The bootstrap resample size \( N \) used in the Bartlett correction is 250 and the number of simulation is 1000.

Tables 1 and 2 contain the empirical coverage and the averaged length of the three types of confidence intervals, which can be summarized as follows. First of all, the need for carrying out the second order correction to the EL confidence interval \( I_\alpha \) is quite obvious as the original EL interval has quite severe under coverage for all the cases considered even for a sample size of 100 for the panel data model. The under coverage is particularly severe when the sample size is small for the normal mean model \( N(1, 2) \) and for the panel data model, These are the situations where the Bartlett correction is needed. It is observed that in all the cases considered the Bartlett correction improves significantly the coverage of \( I_\alpha \). The restoration of coverage by the Bartlett correction is very impressive. We also observed that, as anticipated in (3.3), the direct bootstrap confidence interval has similar performance with the Bartlett corrected intervals in most of the cases. However, in the normal mean models with 90% nominal coverage level, the coverage of the direct bootstrap intervals is not as good as the BC. The robust performance of the Bartlett corrected interval may be due to the fact that the estimation of the Bartlett factor \( \tilde{B}_c \), which involves simple bootstrap averaging, is more robust than the bootstrap estimation of the extreme quantiles of the distribution of \( r(\hat{\theta}) \). We also observed in passing that as the sample size increases the EL interval \( I_\alpha \) improves both in its coverage and length whereas the improvement of the Bartlett intervals is shown in terms of shorter length.

5. Conclusions

The main finding of the paper is that the EL with general moment restrictions are Bartlett correctable. This is a substantial extension of the previously established cases of Bartlett
correction of EL, including the case of smoothed functions of means by DiCiccio, Hall and Romano (1991) more than one decade ago. It shows that the dedicated Bartlett property of the EL is still preserved even in the case of over-identification. Although the Bartlett factor admits a very involved expression with over-identified moment restrictions, proving that the EL is Bartlett correctable in the general case provides the theoretical foundation to the proposed easily implementable empirical Bartlett correction.

The use of the bootstrap to carry out the Bartlett correction empirically is due to a rather involved expression for the Bartlett factor. Although it may be expected that the direct bootstrap calibration would give the same effect as the Bartlett correction, the justification of the direct bootstrap method inevitably needs those cumulants and the Edgeworth expansions established in this paper.

The results established in Theorems 1 and 2 can be extended to independent but not identically distributed samples, for instance those arisen in a regression study. We need to modify $\alpha$, $\beta$ and $\gamma$ as follows:

$$\alpha^{j_1\ldots j_k} = n^{-1} \sum_{i=1}^{n} E[w_i^{j_1}(\theta_0)\ldots w_i^{j_k}(\theta_0)], \quad \beta^{j_1\ldots j_k} = n^{-1} \sum_{i=1}^{n} E\left(\frac{\partial^k \Gamma_i(\theta_0, \theta_0)}{\partial \eta_{j_1} \ldots \partial \eta_{j_k}}\right)$$

$$\gamma^{j_1\ldots j_l; k_1\ldots k_m; p_1\ldots p_n} = \frac{1}{n} \sum_{i=1}^{n} E\left(\frac{\partial^l w_i^{j_1}(\theta_0)}{\partial \theta_{j_1}} \ldots \frac{\partial^m w_i^{j_l}(\theta_0)}{\partial \theta_{j_l}} \ldots \frac{\partial^n w_i^{p_1}(\theta_0)}{\partial \theta_{p_1}} \ldots \frac{\partial^n w_i^{p_n}(\theta_0)}{\partial \theta_{p_n}}\right)$$

We need also to re-define $V_n$ as $n^{-1} \sum_{i=1}^{n} Var\{g(X_i, \theta_0)\}$. These forms of $\alpha$ and $V_n$ were employed in Chen (1993) to establish Bartlett correction for linear regression where $r = p$. The conditions (2.1) should be modified to reflect the independent but not identically distributed nature of data. Similar conditions as those given in Theorem 20.6 of Bhattacharya and Rao (1976) are required, as in Chen (1993, 1994). Then, it may be shown that Theorem 1 is true by employing Skovgaard (1981) on transformation of Edgeworth expansions. Theorem 2 is then a consequence of Theorem 1 as the calculation of the cumulants follows the same spirit given in the Appendix for independent and identically distributed samples.

APPENDIX

We provide some technical details on the log EL ratio $r(\theta_0)$ in A.2, the sign root decom-
position in A.3., and the proofs of the two theorems in A.4.

A.1. Basic formulae.

We first present some basic formulae which will be used throughout the derivations.

Let us define \( \Omega = (\omega^{kl})_{p \times p} =: U \Lambda^{-1} \) where \( \omega^{kl} = u^{kl} \lambda^{-1}_l \). Please note here that no summation over the subscript \( l \) is carried out due to \( \Lambda \) being a diagonal matrix. Since \( \Gamma(\eta) = S^{-1}Q(\eta) \) where

\[
S^{-1} = \begin{pmatrix} -I + S_{12}(S_{12}^T S_{12})^{-1}S_{12}^T & S_{12}(S_{12}^T S_{12})^{-1} \\ (S_{12}^T S_{12})^{-1}S_{12}^T & (S_{12}^T S_{12})^{-1} \end{pmatrix} = \begin{pmatrix} 0 & 0 & \Omega^T \\ 0 & -I_{r-p} & 0 \\ \Omega & 0 & \Omega \Omega^T \end{pmatrix},
\]

it can be checked that

\[
\beta^{j,k} = E(\frac{\partial \Gamma_j(0,\theta_0)}{\partial \eta_k}) = \delta^{j,k} \quad \text{and} \quad B := \begin{pmatrix} B^1 \\ \vdots \\ B^r \end{pmatrix} = S^{-1} \begin{pmatrix} A \\ 0 \end{pmatrix} = \begin{pmatrix} 0 \\ -A_2 \\ \Omega A_1 \end{pmatrix}.
\]

Here \( A^T = (A^1, \ldots, A^r)^T =: (A_1^T, A_2^T)^T \), where \( A_1 = (A^1, \ldots, A^p)^T \) and \( A_2 = (A^{p+1}, \ldots, A^r)^T \) constitute a partition of vector \( A \). Therefore for positive integers \( k \) and \( a \),

\[ (A.1) \quad B^k = 0 \text{ for } k \leq p; \quad B^{p+a} = -A^{p+a} \text{ for } a \leq r - p, \quad \text{and } B^{r+k} = \omega^{kl}A^l \text{ for } k \leq p. \]

Let \( B_1 = (B^1, \ldots, B^r)^T \) and \( B_2 = (B^{r+1}, \ldots, B^{r+p})^T \). Since \( SB = (A^r, 0_{p \times 1})^T \) which means that \( -B_1 + S_{12}B_2 = A \). As \( S_{12} = (\gamma^{j,k})_{r \times p} \) and from (A.1) we have

\[ (A.2) \quad \gamma^{j,k}B^{r+k} = A^j I(j \leq p) \]

where \( I \) is the indicator function. Since

\[ (A.3) \quad (B^{p,a})_{(r+p) \times (r+p)} = S^{-1} \begin{pmatrix} -(A^{ij}) & (C^{ij}) \\ (C^{ij})^T & 0 \end{pmatrix}, \]

we have \( S_{21}(B^{j,k})_{r \times p} = (C^{k,m})_{r \times r} \) and \( S_{21}(B^{j,r+a})_{r \times p} = 0 \). As \( S_{21} = (\gamma^{j,k})^T \), these mean

\[ (A.4) \quad \gamma^{j,k}B^{j,l} = C^{l,k} \text{ for } l \leq r \text{ and } k \leq p \text{ and } \gamma^{j,k}B^{j,r+a} = 0. \]
Furthermore, (A.3) also implies the following which links the $B^{s,t}$ system with the $A^{j,m}$ and the $C^{j,m}$

\[
(A.5) \begin{pmatrix}
(\omega^{mk} C^{l,m}) & (\omega^{mk} C^{p+b,m}) & 0 \\
(A^{p+a,l}) & (A^{p+a,p+b}) & -(C^{p+a,l}) \\
(\omega^{km}[\omega^{nm} C^{l,n} - A^{ml}]) & (\omega^{km}[\omega^{nm} C^{p+b,n} - A^{m p+b}]) & (\omega^{km} C^{t,m})
\end{pmatrix}.
\]

In a similar fashion, we can establish the links between $\beta^{s,t}$ and $(\alpha^j, \gamma^j)$ systems where $t$ and $j$ contains either single or double superscripts. Define $Q_1(\theta, \lambda) = \frac{w(\theta)}{1 + \lambda^T w(\theta)}$ and $Q_2(\theta, \lambda) = \frac{(\partial w(\theta)/\partial \theta)^T \lambda}{1 + \lambda^T w(\theta)}$. For fixed $h \in \{1, \ldots, r\}$ and $k \in \{1, \ldots, p\}$

\[
\frac{\partial Q_1(\theta, \lambda)}{\partial \lambda} = \frac{w(\theta) w^T(\theta)}{(1 + \lambda^T w(\theta))^2}, \quad \frac{\partial Q_1(\theta, \lambda)}{\partial \theta} = \frac{\partial w(\theta)}{\partial \theta} \frac{w^T(\theta)}{(1 + \lambda^T w(\theta))^2},
\]

\[
\frac{\partial}{\partial \lambda} \frac{\partial Q_1(0, \theta_0)}{\partial \lambda} = 2w(\theta_0)w^T(\theta_0)w^h(\theta_0), \quad \frac{\partial}{\partial \lambda} \frac{\partial Q_1(0, \theta_0)}{\partial \theta} = -\frac{\partial w(\theta_0)}{\partial \theta} w^h(\theta_0) - w(\theta_0) \frac{\partial w^h(\theta_0)}{\partial \theta},
\]

\[
\frac{\partial}{\partial \theta} \frac{\partial Q_1(0, \theta_0)}{\partial \theta} = -\frac{\partial w(\theta_0)}{\partial \theta} w^T(\theta_0) - w(\theta_0) \frac{\partial w^T(\theta_0)}{\partial \theta}, \quad \frac{\partial}{\partial \theta} \frac{\partial Q_1(0, \theta_0)}{\partial \theta} = \frac{\partial w(\theta_0)}{\partial \theta} \frac{\partial w^T(\theta_0)}{\partial \theta},
\]

and

\[
\frac{\partial Q_2(\theta, \lambda)}{\partial \lambda} = \frac{\partial w^T(\theta)/\partial \theta}{(1 + \lambda^T w(\theta))^2}, \quad \frac{\partial Q_2(\theta, \lambda)}{\partial \theta} = \frac{\partial w^T(\theta)/\partial \theta}{(1 + \lambda^T w(\theta))^2} - \frac{\partial w^T(\theta)/\partial \theta \partial \theta}{(1 + \lambda^T w(\theta))^2},
\]

\[
\frac{\partial}{\partial \lambda} \frac{\partial Q_2(0, \theta_0)}{\partial \lambda} = -\frac{\partial w^T(\theta_0)}{\partial \theta} w^h(\theta_0) - \frac{\partial w^h(\theta_0)}{\partial \theta} w^T(\theta_0), \quad \frac{\partial}{\partial \lambda} \frac{\partial Q_2(0, \theta_0)}{\partial \theta} = \frac{\partial^2 w^h(\theta_0)}{\partial \theta \partial \theta},
\]

\[
\frac{\partial}{\partial \theta} \frac{\partial Q_2(0, \theta_0)}{\partial \theta} = -\frac{\partial w^T(\theta_0)}{\partial \theta} w^h(\theta_0) - \frac{\partial w^h(\theta_0)}{\partial \theta} w^T(\theta_0), \quad \frac{\partial}{\partial \theta} \frac{\partial Q_2(0, \theta_0)}{\partial \theta} = \frac{\partial^2 w^h(\theta_0)}{\partial \theta \partial \theta} = 0.
\]

These mean that for a fixed $h \in \{1, \ldots, r\}$

\[
(\beta^{s,t,h}) = \begin{pmatrix}
0_{p \times p} & 0 & \Omega^T \\
0 & -I_{r-p} & 0 \\
\Omega & 0 & \Omega \Omega^T
\end{pmatrix} \times E \begin{pmatrix}
\frac{\partial^2 Q_1}{\partial \lambda \partial \lambda} & \frac{\partial^2 Q_1}{\partial \lambda \partial \theta} \\
\frac{\partial^2 Q_2}{\partial \lambda \partial \lambda} & \frac{\partial^2 Q_2}{\partial \lambda \partial \theta}
\end{pmatrix}
\]
\[
\begin{pmatrix}
0_{p \times p} & 0 & \Omega^T \\
0 & -I_{r-p} & 0 \\
\Omega & 0 & \Omega \Omega^T
\end{pmatrix} \times 
\begin{pmatrix}
(2\alpha^{klh})_{p \times p} & (2\alpha^k p+b, h)_{p \times r-p} & - (\gamma^{h,k,m} + \gamma^{k,h,m})_{p \times p} \\
(2\alpha^{p+a, lh})_{r-p \times p} & (2\alpha^{p+a, p+b, h})_{r \times r-p} & - (\gamma^{h,p+a,m} + \gamma^{p+a,h,m})_{r \times p} \\
- (\gamma^{h,k,m} + \gamma^{k,h,m})_{p \times p} & - (\gamma^{h,p+a,m} + \gamma^{p+a,h,m})_{p \times r-p} & (\gamma^{h,k,l})_{r \times p}
\end{pmatrix}
\]

and for a fixed \( h = r + k \) where \( k \in \{1, \cdots, p\} \)
\[
(\beta^{s,tr+k}) = 
\begin{pmatrix}
0 & 0 & \Omega^T \\
0 & -I_{r-p} & 0 \\
\Omega & 0 & \Omega \Omega^T
\end{pmatrix} \times 
\begin{pmatrix}
\frac{\partial^2 Q_1}{\partial x^2} & \frac{\partial^2 Q_2}{\partial y^2} \\
\frac{\partial^2 Q_3}{\partial y \partial x} & \frac{\partial^2 Q_4}{\partial x \partial y}
\end{pmatrix}
\]

\[
= 
\begin{pmatrix}
0 & 0 & \Omega^T \\
0 & -I_{r-p} & 0 \\
\Omega & 0 & \Omega \Omega^T
\end{pmatrix} \times 
\begin{pmatrix}
\gamma^{l,k,m} + \gamma^{l,m,k} & \gamma^{l,k,p+a} + \gamma^{k,p+a,k} & (\gamma^{l,mk})_{p \times p} \\
- (\gamma^{p+a,k,m} + \gamma^{p+a,m,k})_{p \times p} & - (\gamma^{p+a,k,p+b} + \gamma^{p+a,p+b,k})_{r \times p} & (\gamma^{p+a, mk})_{r \times p} \\
(\gamma^{l,mk})_{p \times p} & (\gamma^{p+a, mk})_{p \times r-p} & 0_{p \times p}
\end{pmatrix}
\]

where \( k, l, m \in \{1, \cdots, p\} \) and \( a, b \in \{1, \cdots, r-p\} \). Hence we have

\[(A.6)
\begin{align*}
\beta^{l,p+a, p+c} &= -\omega^d [\gamma^{p+c+p+a, o} + \gamma^{p+a+p+c, o}], \\
\beta^{l,p+m, p+c} &= \omega^d [\gamma^{p+c+o, m}], \\
\beta^{p+a,p+b, p+c} &= -2\alpha^{p+a, p+b, p+c}, \\
\beta^{p+a,p+m, p+c} &= \gamma^{p+c+p+a, m} + \gamma^{p+a+p+c, m}, \\
\beta^{l,p+a, p+n} &= \omega^d [\gamma^{p+a, p+n}], \\
\beta^{l,p+m, p+n} &= 0, \\
\beta^{p+a,p+b, p+n} &= \gamma^{p+a, p+b+n} + \gamma^{p+a+p+b, n}, \\
\beta^{p+a,p+m, p+n} &= -\gamma^{p+a, p+m+n}, \\
\beta^{r+k,p+a, p+c} &= 2\omega^k \alpha^{p+a, p+c} - \omega^k \omega^a \gamma^{p+c+p+a, n} + \gamma^{p+a+p+c, n}, \\
\beta^{r+k,p+a, p+n} &= \omega^k \omega^a \gamma^{p+a, p+n} - \omega^k [\gamma^{o, p+a, n} + \gamma^{o, p+a, n}], \\
\beta^{r+k,p+m, p+c} &= \omega^k \omega^a \gamma^{p+c+m} - \omega^k [\gamma^{o, p+c+m} + \gamma^{o, p+c+m}], \\
\beta^{r+k,p+m, p+n} &= \omega^k [\gamma^{o, p+c+m}].
\end{align*}
\]

A.2. Derivations of (2.6), (2.8) and (2.9)
Derivation of (2.6). As shown in Appendix 1, \( \beta^{jk} = \delta^{jk} \). Now expanding \( \Gamma(\hat{\theta}, \hat{\lambda}) \) around \((0, \theta_0)\),

\[
0 = \Gamma^j(\hat{\theta}, \hat{\lambda}) = B^j + \beta^{jk}(\hat{\eta}^k - \eta_0^k) + B^{jk}(\hat{\eta}^k - \eta_0^k)
+ \frac{1}{2} \beta^{j,kl}(\hat{\eta}^k - \eta_0^k)(\hat{\eta}^l - \eta_0^l) + \frac{1}{3} B^{j,kl}(\hat{\eta}^k - \eta_0^k)(\hat{\eta}^l - \eta_0^l)(\hat{\eta}^m - \eta_0^m)
+ \frac{1}{6} \beta^{j,klm}(\hat{\eta}^k - \eta_0^k)(\hat{\eta}^l - \eta_0^l)(\hat{\eta}^m - \eta_0^m)(\hat{\eta}^n - \eta_0^n)
+ \frac{1}{24} B^{j,klmn}(\hat{\eta}^k - \eta_0^k)(\hat{\eta}^l - \eta_0^l)(\hat{\eta}^m - \eta_0^m)(\hat{\eta}^n - \eta_0^n) + O_p(n^{-3}).
\]

Converting the above expansion,

\[
\hat{\eta}^j - \eta_0^j = -B^j + B^{j,k}B^k - B^{j,k}B^{k,l}B^l + \frac{1}{2} \beta^{j,lm}B^{j,k}B^l B^m
- \frac{1}{6} \beta^{j,kl}[B^{k,l}B^l - B^{k,m}B^m B^l - B^{l,q}B^q B^k + \frac{1}{2} \beta^{k,lm}B^m B^n B^l + \frac{1}{2} \beta^{l,qs}B^q B^s B^k]
- \frac{1}{2} B^{j,kl}B^k B^l + \frac{1}{6} \beta^{j,klm}B^k B^l B^m + O_p(n^{-2})
= -B^j + B^{j,k}B^k - \frac{1}{2} \beta^{j,kl}B^k B^l - B^{j,k}B^{k,l}B^l + \frac{1}{2} \beta^{j,lm}B^{j,k}B^l B^m + \beta^{j,kl}B^{k,m}B^m B^l
- \frac{1}{2} \beta^{j,kl}B^{k,m}B^m B^n B^l - \frac{1}{2} B^{j,kl}B^{k,l}B^l + \frac{1}{6} \beta^{j,klm}B^k B^l B^m + O_p(n^{-2}),
\]

where \( j, k, l, m, \in \{1, 2, \cdots, r + p\} \). Thus, we have established (2.6).

From (2.6), we have the following expansions for \( \hat{\lambda} \) and \( \hat{\theta} \):

\[
\hat{\lambda}^j = -B^j + B^{j,a}B^a - \frac{1}{2} \beta^{j,au}B^u B^a - B^{j,u}B^u B^a + \frac{1}{2} \beta^{j,as}B^a B^s B^a + \frac{1}{6} \beta^{j,ags}B^a B^s B^a B^a
+ \frac{1}{2} \beta^{j,au}B^u B^a + \frac{1}{6} \beta^{j,ags}B^a B^s B^a B^a + O_p(n^{-2}),
\]

(A.7)

and

\[
\hat{\theta}^k = -B^{r + k} + B^{r + k,u}B^u B^a - \frac{1}{2} \beta^{r + k,au}B^a B^u B^a - B^{r + k,u}B^{u,a}B^a
+ \frac{1}{2} \beta^{u,as}B^{r + k,u}B^a B^s + \beta^{r + k,u}B^u B^u B^s - \frac{1}{2} \beta^{u,as}B^{r + k,u}B^a B^s B^a
+ \frac{1}{6} \beta^{u,ags}B^a B^s B^a B^a + O_p(n^{-2}).
\]

(A.8)

Derivation of (2.8) and (2.9). We shall expand each term on the right of (2.7). By
ignoring terms of $O_p(n^{-5/2})$, the first term

$$
\hat{\lambda}^T n^{-1} \sum_{i=1}^n w_i(\hat{\theta}) = \hat{\lambda}^T n^{-1} \sum_{i=1}^n \left[ w_i^2(\theta_0) + \frac{\partial w_i^2(\theta_0)}{\partial \hat{\theta}} \hat{\theta}^k + \frac{1}{2} \frac{\partial^2 w_i^2(\theta_0)}{\partial \hat{\theta} \partial \hat{\theta}} \hat{\theta}^k \hat{\theta}^l + \frac{1}{6} \frac{\partial^3 w_i^2(\theta_0)}{\partial \hat{\theta} \partial \hat{\theta} \partial \hat{\theta}} \hat{\theta}^k \hat{\theta}^l \hat{\theta}^m \right] \\
= \hat{\lambda}^T A^i + \gamma_{j,k} \hat{\lambda}^T \hat{\theta}^k + \frac{3}{2} \gamma_{j,k} \hat{\lambda}^T \hat{\theta}^k \hat{\theta}^l + \frac{3}{2} \gamma_{j,k} \hat{\lambda}^T \hat{\theta}^k \hat{\theta}^l \hat{\theta}^m.
$$

Similarly, the second term

$$
\hat{\lambda}^T n^{-1} \sum_{i=1}^n w_i(\hat{\theta}) w_i(\hat{\theta})^T \hat{\lambda} \\
= \hat{\lambda}^T \hat{\lambda}^h n^{-1} \sum_{i=1}^n \left\{ w_i^2(\theta_0) w_i^2(\theta_0) + w_i^2(\theta_0) \frac{\partial w_i^2(\theta_0)}{\partial \hat{\theta}} \hat{\theta}^i [2, j, h] + \frac{1}{2} w_i^2(\theta_0) \frac{\partial^2 w_i^2(\theta_0)}{\partial \hat{\theta} \partial \hat{\theta}} \hat{\theta}^i \hat{\theta}^j \right\} \\
= \hat{\lambda}^T \hat{\lambda}^i (A_{ji} + \delta_{ji}) + \frac{1}{2} \hat{\lambda}^T \hat{\lambda}^i \hat{\theta}^j \{ (C_{ji,k} + \gamma_{j,k})[2, j, i] \} + \frac{1}{2} \hat{\lambda}^T \hat{\lambda}^i \hat{\theta}^j \{ (C_{ji,k} + \gamma_{j,k})[2, j, i] \} \\
+ \frac{1}{2} \hat{\lambda}^T \hat{\lambda}^i \hat{\theta}^j \{ (C_{ji,k} + \gamma_{j,k}) \} \\
= \hat{\lambda}^T \hat{\lambda}^i + \frac{1}{2} \hat{\lambda}^T \hat{\lambda}^i \hat{\theta}^j + 2 \gamma_{j,i,k} \hat{\lambda}^T \hat{\theta}^j + 2 \gamma_{j,i,k} \hat{\lambda}^T \hat{\theta}^j + (\gamma_{j,i,k} + \gamma_{j,i,k}) \hat{\lambda}^T \hat{\theta}^j + O_p(n^{-5/2})
$$

where $[2, j, i]$ indicates there are two terms by swamping the super-scripts $i$ and $j$, and the same is understood for similar notations. For the third term

$$
\frac{2}{3} n^{-1} \sum_{i=1}^n [\hat{\lambda}^T w_i(\hat{\theta})]^3 = \frac{2}{3} \hat{\lambda}^T \hat{\lambda}^h (A_{jih} + \alpha_{jih}) + \frac{2}{3} \{ \hat{\lambda}^T \hat{\lambda}^i \hat{\lambda}^j \hat{\theta}^k \} (3, j, i, h) \\
= \frac{2}{3} \hat{\lambda}^T \hat{\lambda}^h (A_{jih} + \alpha_{jih}) + 2 \hat{\lambda}^T \hat{\lambda}^i \hat{\theta}^j \hat{\theta}^k \hat{\theta}^l + O_p(n^{-5/2}).
$$

Finally, $n^{-1} \sum_{i=1}^n [\hat{\lambda}^T w_i(\hat{\theta})]^4 = \hat{\lambda}^T \hat{\lambda}^h \hat{\lambda}^i \hat{\theta}^j \hat{\theta}^k + O_p(n^{-5/2})$.

We then have for $a, b, c, d \in \{1, 2, \ldots, r - p\}$, $f, g, h, i, j \in \{1, 2, \ldots, r\}$, $k, l, m, n, o \in \{1, 2, \ldots, p\}$ and $q, s, t, u \in \{1, 2, \ldots, r + p\}$ that

$$
n^{-1} l(\hat{\theta}) = -2B^i A^j - B^i B^j + 2B^{i,a} B^a (A^j + B^j) - \beta_{j,a} B^a B^q (A^j + B^j) \\
- 2B^{i,a} B^a (A^j + B^j) + \beta_{j,a} B^j B^a B^q (A^j + B^j) - \beta_{j,a} B^a B^q (A^j + B^j) \\
- B^{j,a} B^a (A^j + B^j) + \frac{1}{2} \beta_{j,a} B^a B^q (A^j + B^j) + 2\beta_{j,a} B^a B^q (A^j + B^j) \\
+ \gamma_{j,k} \{ -B^{j,a} B^q B^{r,k} + \frac{1}{2} \beta_{j,a} B^a B^q B^{r,k} + B^{j,a} B^q B^{r,k} \}.
$$
Applying (A.1) and (A.4), it may be shown that the 3rd to the 18th terms on the right hand side cancel each other and the application of (A.4) simplifies the 20th term. Keep all the

Now bring in the expansion for \( \ell(\theta_0) \) in (2.5) we have

\[
\begin{align*}
n^{-1}r(\theta_0) &= (A^j + B^j)(A^j + B^j) - A^{ji}(A^j A^i - B^j B^i) - 2C^{ji,k} B^j B^{r+k} - 2\gamma^{ji,l} B^j B^i B^{r+l} \\
&+ \frac{2}{3} \alpha^{jih}[A^{j}A^{h}B^{j}B^{h}] + \gamma^{j,kl} B^j B^{r+k} B^{r+l} + A^{ji}(A^{hi}A^{jh} - B^{ij}B^{hi}B^{j}[2,i,j]) \\
&- \beta^{j,uq}(C^{j,k} B^{r+k} - B^{r+k,s} B^s \gamma^{j,k} + B^{j,s} B^s - A^{ji} B^i) B^u B^a - 2\alpha^{jih} B^j B^i B^{h,q} B^q \\
&+ B^{iu} B^{j,q} B^u B^a + 2C^{ji,k} B^{j,a} B^{r+k} B^q - \gamma^{j,kl} B^j B^{r+k} B^{r+l} B^{j,a} B^q \\
&- 2\gamma^{j,kl} B^j B^{r+l} B^{r+k,a} B^q + 2\gamma^{j,i,l}(B^j B^{r+l,q} B^q + B^{r+l} B^{j,l} B^q B^h B^i B^{r+k} \\
&+ \left( \frac{1}{2} \beta^{j,uq} \beta^{j,sl} - \frac{1}{2} \beta^{j,uq} \beta^{r+k,s} \gamma^{j,k} \right) B^u B^q B^s B^t + \frac{1}{2} \beta^{j,uq} B^u B^q B^{r+l} B^{r+k} \\
&+ (\gamma^{j,l} B^{r+l} B^q - \gamma^{j,l} B^{r+l} B^{r+k} - 2\gamma^{j,i,h,k} B^j B^i B^h B^{r+k} \\
&+ (\gamma^{j,i,k} + \gamma^{j,i,k}) B^j B^i B^{r+l} B^{r+k} - \frac{1}{3} \gamma^{j,kml} B^j B^{r+k} B^{r+l} B^{r+m}
\end{align*}
\]
+ \frac{1}{2}\alpha^{ijh}B^jB^iB^kB^q + (\alpha^{ijh}\beta^{l,uj} - \gamma^{j,l,j^{r+l,uj}})B^jB^iB^uB^q
\]
\]
\begin{align*}
+ C_j^{,kl}B^jB^{r+k}B^{r+l} - 2C_j^{,lj}B^jB^iB^{r+l} + \frac{2}{3}A^{ijh}(B^jB^kB^h + A^jA^iA^h) \\
(2.9)
- 2\alpha^{ijh}A^g A^j A^i A^g + \alpha^{jg,\ell} A^j A^i A^h A^g - \frac{1}{2}\alpha^{ijh}A^j A^i A^h A^g + O_p(n^{-5/2}).
\end{align*}

Now the terms appeared on the 3rd line of the above equation cancel each other. To appropriate this, by applying the relationships implied by (A.5) together with (A.1) and (A.4),

\[
\begin{align*}
\beta^{i,j,\ell}[C_j^{,kl}B^{r+k} - B^{r+k,s}B^{s,j,k} + B^j B^i B^a - A^{i,j} B^a]B^a B^q \\

= [\beta^{i,j,\ell}C_j^{,kl}B^{r+k} - \beta^{j,\ell,l}B^{r+k,l} + \beta^{i,\ell}B^{s,j,k} + \beta^{i,\ell}B^j B^i B^a - \beta^{i,\ell}B^{a,j,k}B^r + B^{r+k,s}B^{s,j,k} + B^j B^i B^a - A^{i,j} B^a]B^a B^q \\

- \beta^{i,\ell}B^p + \beta^{i}B^{p,a} + \beta^{i,\ell}B^{p,a}B^p + \beta^{i,\ell}B^{p,a}B^{p,a} + \beta^{i,\ell}B^{p,a}B^{p,a} + \beta^{i,\ell}B^{p,a}B^{p,a} + \beta^{i,\ell}B^{p,a}B^{p,a} \\

+ \beta^{i,\ell}B^p + \beta^{i}B^{p,a} + \beta^{i,\ell}B^{p,a}B^p + \beta^{i,\ell}B^{p,a}B^{p,a} + \beta^{i,\ell}B^{p,a}B^{p,a} + \beta^{i,\ell}B^{p,a}B^{p,a} \\

- \beta^{i,\ell}B^p + \beta^{i}B^{p,a} + \beta^{i,\ell}B^{p,a}B^p + \beta^{i,\ell}B^{p,a}B^{p,a} + \beta^{i,\ell}B^{p,a}B^{p,a} + \beta^{i,\ell}B^{p,a}B^{p,a} \\

\end{align*}
\]

Applying again (A.5), we can express the terms appeared in the first two lines of (A.9) in term of As, that is

\[
(A^j + B^j)(A^i + B^i) - A^{ij}(A^j A^i - B^j B^i) - 2C_j^{,kl}B^j B^{r+k} \\
- \gamma^{j,l,l}B^j B^r + [2, j, i] + \frac{\alpha}{\beta}C_{j,kl}B^{r+k}B^{r+l} \\
= A^j A^l - (A^k A^l A^k + 2A^j A^l A^p + A^p A^a A^l) + 2\omega^{kl}C_{p,a,k}A^{p,a}A^l \\
+ (\gamma^{p,a,b,k} + \gamma^{p,b,p+a,k})\omega^{kl}A^{p,a}A^{p+a}A^l + \frac{2}{\beta}A^{k}A^{m}A^{k}A^{m} \\
+ 2\omega^{kl}C_{p,a,k}A^{p,a}A^l + 2\omega^{kl}C_{p,a,k}A^{p,a}A^l + 2\omega^{kl}C_{p,a,k}A^{p,a}A^l \\
+ 2\omega^{kl}C_{p,a,k}A^{p,a}A^l + 2\omega^{kl}C_{p,a,k}A^{p,a}A^l \\
+ 2\omega^{kl}C_{p,a,k}A^{p,a}A^l.
\]

Substitute these results into (2.9), we arrive at (2.9).
A.3. Expansion for $R_3$

We subtract $R_2^l R_2^l$ from all the terms appeared in line 4 and below in (2.9). Fortunately all the terms which do not have $A^l$ appeared cancel out with those appeared in $R_2^l R_2^l$. Otherwise, a signed root decomposition of the EL ratio $r(\theta_0)$ would not be possible. Hence the remaining terms can be written as $2R_1^l R_3^l$.

The pursuit for an expression of $R_3$ is done by repeatedly employing the formulae (A.5) and (A.6) as well as (A.1), (A.2) and (A.4). For instance, the terms appeared in the 4th line of (2.9)

$$A^j i A^h A^i A^h + B^{i,u} B^{j,q} B^u B^q - 2A^{ij} B^{i,q} B^q B^j + 2C^{j,k} B^{j,q} B^{r+k} B^q$$

$$= A^j i A^h A^m A^l + 2A^{jl} A^k p+a A^p+a A^l + 2A^l p+a A^p+a p+b A^p+b A^l$$

$$= A^j p+a A^i p+a A^k A^l + A^p+a A^l A^p+a A^p+b + \omega^{ml} C^p+a,m C^p+b,n A^p+a A^p+b$$

$$+ \omega^{nl} C^p+a,m C^p+a,k A^m A^l - 2\omega^{mk} C^p+b,m A^p+a k A^p+b A^l$$

$$- 2\omega^{ml} \omega^{kn} C^p+a,k C^p+a,m A^m A^l - 2\omega^n C^n,k C^p+a,k A^p+a A^l$$

$$- 2\omega^{ml} \omega^{kn} C^p+a,k C^p+a,m A^m A^l$$

and the terms in the fifth line

$$-\gamma^{jk,kl} B^{r+k} B^{r+l} B^{j,q} B^q - 2\gamma^{j,kl} B^j B^l B^{r+k} B^q$$

$$= -\gamma^{ml} C^p+a B^{r+l} B^{p+a} B^{m,p+a} - \gamma^{p+b,kl} B^p+a B^{r+k} B^{r+l} B^p p+a$$

$$= -\gamma^{p+b,kl} B^{r+k} B^{r+l} B^{r+n} B^{b,r+n} - 2\gamma^{p+b,kl} B^p+a B^p+b B^{p+b}$$

$$- 2\gamma^{p+a,kl} B^p+a B^{r+l} B^{r+m} B^{r+k} r+m$$

$$= \gamma^{ml} B^{r+l} B^{r+m} B^{r+k} r+m$$

and so on for the other terms. It may be shown after some quite involved algebra that that

$$R_3^l = \sum_{i=1}^{6} R_{3i}^l,$$

where

$$R_{3i}^l = \frac{3}{8} A^i A^m A^k + \frac{1}{3} A^i A^k A^m - \frac{5}{12} A^i A^m A^k A^n$$

$$= 21.$$
\[
\begin{align*}
- \frac{5}{12} \alpha^{k,mn} A^{lm} A^k A^n + \frac{4}{9} \alpha^{lkn} \alpha^{mnn} A^m A^k A^o - \frac{1}{4} \alpha^{lkm} A^m A^k A^n, \\
\mathcal{P}_{32}^l &= A^{lk} p+a A^{lp+a} A^k + A^{l} p+a p+b A^{lp+a} A^{lp+b} - \frac{1}{2} \omega^{lk} \omega^m C^{p+a,k} A^{lp+a} A^n \\
&\quad - \omega^{lk} C^{p+ap+b,k} A^{lp+a} A^{lp+b} + \frac{1}{2} \omega^{lk} C^{p+a,k} A^{lm} A^{lp+a} + \frac{1}{2} A^{lk} A^k p+a A^{lp+a} \\
&\quad + A^{lp+a} A^{lp+a} p+b A^{lp+b} + \frac{1}{2} A^{lp+a} A^{lp+b} A^k - \frac{1}{2} \omega^{lk} C^{p+a,n} C^{p+a,k} A^m \\
&\quad - \omega^{lk} \omega^{mn} C^{n,k} C^{p+a,m} A^{lp+a} - \omega^{lk} C^{p+ap+b,k} A^{lp+a} p+b A^{lp+b}, \\
\mathcal{P}_{33}^l &= \frac{1}{2} \gamma^{m,rn} \omega^m \omega^l \omega^k - \frac{1}{3} \alpha^{mn} \omega^{km} C^{p+ak} A^{lp+a} A^n + \frac{1}{2} \gamma^{lp+ak} A^{lp+a} A^v + \omega^{vl} \omega^{k,l} (\gamma^{lp+ak} + \gamma^{lp+ak}) \\
&\quad - \alpha^{lp+k} + \frac{1}{2} \gamma^{lp+ak} \omega^m C^{p+ak} A^{lp+a} A^n + \gamma^{lp+a,k} \omega^m A^{lp+a} A^n + \frac{1}{2} \gamma^{lp+ak} A^{lp+a} A^v \\
&\quad - \frac{1}{2} \gamma^{lp+ak} A^{lp+a} A^{lp+b}, \\
\mathcal{P}_{34}^l &= \gamma^{p+ap+b} A^{lp+a} A^{lp+b} - \alpha^{lp+a} p+b A^{lp+a} A^{lp+b} + \frac{1}{2} (\gamma^{lp+ap+b} \omega^m C^{p+ak} A^{lp+a} A^n - \alpha^{lp+a} A^m A^v) \\
&\quad + \left( \gamma^{lp+ap+b} A^{lp+a} A^v \right) \omega^{mn} A^{lp+a} A^n - \alpha^{lp+a} A^m A^v \\
&\quad - \frac{1}{2} \alpha^{lp+ap+b} A^{lp+a} A^{lp+b}, \\
\mathcal{P}_{35}^l &= \frac{1}{2} \alpha^{lp+ap+b} A^{lp+a} A^{lp+b} - \frac{1}{8} \omega^{lm} \omega^{m,\omega^{n',\omega^{k',\omega^{p+a,m',\omega^{n',\omega^{l',\omega^{p+a,n'k'}}}}} A^m A^n A^k} \\
&\quad + \frac{2}{3} \alpha^{lp+ap+b} A^{lp+a} A^{lp+b} - \frac{1}{3} \alpha^{lp+ap+b} A^{lp+a} A^n \\
&\quad - \frac{1}{3} \alpha^{lp+ap+b} A^{lp+a} A^n \\
&\quad + \left( \alpha^{lp+ap+b} + \alpha^{lp+ap+b} \omega^{mn} A^{lp+a} A^n \right) \omega^{mn} A^{lp+a} A^n \\
&\quad + \left\{ \alpha^{lp+ap+b} + \alpha^{lp+ap+b} \omega^{mn} A^{lp+a} A^n \right\} \omega^{mn} A^{lp+a} A^n \\
&\quad \times \left( \alpha^{lp+ap+b} + \alpha^{lp+ap+b} \omega^{mn} A^{lp+a} A^n \right) \omega^{mn} A^{lp+a} A^n \\
&\quad \times \left( \alpha^{lp+ap+b} + \alpha^{lp+ap+b} \omega^{mn} A^{lp+a} A^n \right) \omega^{mn} A^{lp+a} A^n \\
\mathcal{P}_{36}^l &= \left\{ \alpha^{lp+ap+b} + \alpha^{lp+ap+b} \omega^{mn} A^{lp+a} A^n \right\} \omega^{mn} A^{lp+a} A^n \\
&\quad - \frac{1}{2} \left( \gamma^{lp+ap+b,m'} + \gamma^{lp+ap+b,m'} \right) \left( \gamma^{lp+ap+b,m'} + \gamma^{lp+ap+b,m'} \right)
\end{align*}
\]
A.4. Proof of Theorems 1 and 2.

The proof of Theorem 1 is divided into two parts. In the first part, we derive the cumulants of $\sqrt{n}R$. In the second part, we establish an Edgeworth expansion for the signed root which then leads to an Edgeworth expansion for the EL ratio $r(\theta_0)$.

Cumulants of the signed root $R$. Since the cumulants of order higher than four are of $O(n^{-2})$ or smaller, we only need to derive the first four cumulants. As the first and the third cumulants are easier to derive than the second and the fourth, we present them first. From (2.10) and (2.11), and the fact that $R^3$ is the product of four zero-mean averages, we have

$$E(R^1_l) = 0, \ E(R^2_l) = n^{-1} \mu^l \text{ and } E(R^3_l) = O(n^{-2})$$

where $\mu^l = -\frac{1}{6} n^{-1} \alpha^{lkk}$. Therefore, the first order cumulant is

$$(A.10) \quad \text{cum}(R^l) = n^{-1} \mu^l + O(n^{-2}).$$

The joint third-order cumulants

$$\text{cum}(R^l, R^o, R^v) = E(R^l R^o R^v) - E(R^l)E(R^o)E(R^v)[3] + 2E(R^l)E(R^o)E(R^v)[3]$$

$$= E(R^l_1 R^o_1 R^v_1) + E(R^l_2 R^o_1 R^v_1)[3] - E(R^l_2)E(R^o_1)E(R^v_1)[3] + O(n^{-3}).$$

We note that

$$E(R^l_1 R^o_1) = n^{-1} \delta^{lo} \quad \text{and}$$

$$E(R^l_2 R^o_1) = n^{-2} [-\frac{1}{2}( \alpha^{lkk} - \delta^{lo}) - \alpha^{lo} p+a p+a + \frac{1}{3} \alpha^{lkm} \alpha^{okm} + \omega^{kl} \gamma^{p+a; a;p+b,k}$$

$$+ (\alpha^{lk} p+a - \frac{1}{2} \omega^{mk} \omega^{nl} \gamma^{p+a, mn}) \alpha^{ok} p+a + (\alpha^{l} p+a p+b - \omega^{kl} \gamma^{p+a; p+b,k}) \alpha^{o} p+a p+b].$$

Write $R^l_2 = R^l_{21} + R^l_{22}$ where $R^l_{21} = -\frac{1}{2} A^{kl} A^k + \frac{1}{3} \alpha^{klm} A^k A^m$ and $R^l_{22}$ contains the rest of the terms appeared in (2.11). We have

$$E(R^l_{21}) = -\frac{1}{6} n^{-1} \alpha^{lkk},$$

23
\[ E(R_{12}^l R_{13}^l) = E(R_{22}^l) \delta^{ow} + O(n^{-3}) = O(n^{-3}) \] and
\[ E(R_{12}^l R_{14}^l) = n^{-2}(-\frac{1}{6} \alpha^{kk} \delta^{ow} - \frac{1}{3} \alpha^{low}) + O(n^{-3}). \]

Thus,
\[ E(R_{21}^l R_{14}^l) = E(R_{21}^l) E(R_{14}^l R_{14}^l) - \frac{1}{3} E(R_{14}^l R_{14}^l R_{14}^l) + O(n^{-3}), \]

which means that
\[ \text{cum}(R_1^l, R_2^l, R_3^l) = O(n^{-3}). \]

To compute the second cumulants, we have to derive the expectation involving the following 21 terms in \( R_1^l R_2^l; \)

\begin{enumerate}
  \item \( \frac{1}{4} A^{kl} A^{kl} A^{ml} A^{m}; \)
  \item \( A^l p+a A^l p+b A^p+a A^p+b; \)
  \item \( \frac{1}{3} \alpha^{klm} \alpha^{nlo} A^k A^m A^n A^o; \)
  \item \( \omega^{kl} \omega^{ml} C^{p+a} k C^{p+b} m A^{p+a} A^{p+b}; \)
  \item \( [\alpha^{kl} p+a - \frac{1}{2} \gamma^{p+a, mn \omega, m} \omega^{n}][\alpha^{ol} p+b - \frac{1}{2} \gamma^{p+b, m' o \omega, n'}] A^{p+a} A^{p+b} A^k A^o; \)
  \item \( [\alpha^l p+a p+b - \gamma^{p+a, p+b, k, \omega, kl}][\alpha^l p+c p+d - \gamma^{p+c, p+d, o \omega, ol}] A^{p+a} A^{p+b} A^{p+c} A^{p+d}; \)
  \item \( A^{kl} A^l p+a A^p+a A^k; \)
  \item \( -\frac{1}{3} \alpha^{klm} A^{ml} A^k A^m; \)
  \item \( -\omega^{kl} C^{p+a} k A^{ml} A^{p+a} A^m; \)
  \item \( -[\alpha^{kl} p+a - \frac{1}{2} \gamma^{p+a, mn \omega, m} \omega^{n}]} A^{p+a} A^{al} A^k A^o; \)
  \item \( -[\alpha^l p+a p+b - \gamma^{p+a, p+b, k, \omega, kl}] A^{ml} A^{p+a} A^{p+b} A^m; \)
  \item \( -[\alpha^l p+a p+b - \gamma^{p+a, p+b, k, \omega, kl}] A^{p+c} A^{p+a} A^{p+b}; \)
  \item \( -2 \omega^{kl} C^{p+a} k A^{p+b} A^{p+a} A^{p+b}; \)
  \item \( -2[\alpha^{kl} p+a - \frac{1}{2} \gamma^{p+a, mn \omega, m} \omega^{n}]} A^{p+b} A^{p+a} A^k; \)
  \item \( -2[\alpha^l p+a p+b - \gamma^{p+a, p+b, k, \omega, kl}] A^{p+c} A^{p+a} A^{p+b}; \)
  \item \( \frac{1}{3} \alpha^{nlm} \omega^{kl} C^{p+a} k A^{p+a} A^n A^m; \)
  \item \( \frac{2}{3} \alpha^{nlm} \omega^{kl} C^{p+a} k A^{p+a} A^n A^m; \)
  \item \( \frac{1}{3} \alpha^{nlm} \omega^{kl} C^{p+a} k A^{p+a} A^n A^m; \)
  \item \( \frac{2}{3} \alpha^{nlm} \omega^{kl} C^{p+a} k A^{p+a} A^n A^m; \)
  \item \( 2 \omega^{al} [\alpha^{kl} p+a - \frac{1}{2} \gamma^{p+a, mn \omega, m} \omega^{n}]} A^{p+a} A^k A^o A^o; \)
  \item \( 2 \omega^{al} [\alpha^l p+a p+b - \gamma^{p+a, p+b, n, \omega, ol}] A^{p+a} A^{p+b} A^k A^o; \)
  \item \( 2 \omega^{al} [\alpha^{kl} p+a - \frac{1}{2} \gamma^{p+a, mn \omega, m} \omega^{n}]} C^{p+b} o A^{p+b} A^p A^k; \)
  \item \( 2 \omega^{al} [\alpha^l p+a p+b - \gamma^{p+a, p+b, k, \omega, kl}] C^{p+c} o A^{p+c} A^p A^k; \)
  \item \( 2 \omega^{al} [\alpha^{kl} p+a - \frac{1}{2} \gamma^{p+a, mn \omega, m} \omega^{n}]} [\alpha^l p+a p+b - \gamma^{p+a, p+b, o \omega, ol}] A^{p+a} A^{p+b} A^p A^k. \)
\end{enumerate}
Denote "[2]" for the permutation of l and o, the expectations for the above 21 terms are respectively

(1) \( \frac{4}{3}[\alpha lkk\alpha omm + \alpha lkm\alpha okm + \alpha lkk - \delta l{o}] \);
(2) \( \alpha l{o} p+a p+a + \alpha l p+a p+a\alpha o p+t b p+b + \alpha l p+a p+b\alpha o p+a p+b \);
(3) \( \frac{4}{3}[\alpha lkk\alpha omm + 2\alpha lkm\alpha okm] \);
(4) \( \omega^{kl}\omega^{mo}[\gamma p+a k p+a m + \gamma p+a k p+a b p+ b + \gamma p+a k p+b m p+a] \);
(5) \( [\alpha l p+a \omega^{kl} + \frac{1}{2} \gamma p+a m n k\omega^{nkl}][\alpha k o p+a - \frac{1}{2} \gamma p+a m n^l k\omega^{nkl} \] [\alpha o p+a p+b - \gamma p+a p+b, \omega^{vo}]\);
(6) \( [\alpha l p+a p+a - \gamma p+a p+a k\omega^{kl}][\alpha o p+a p+b - \gamma p+a p+b, \omega^{vo}]\);
(7) \( \frac{2}{3}[\alpha lkk\alpha o p+a p+a + \alpha l k p+a o p+a\alpha o p+a \alpha o p+a] \); (8) \(-\frac{1}{3}[\alpha lkk\alpha omm + 2\alpha lkm\alpha okm] \);
(9) \(-\frac{1}{3}[\alpha l p+a p+a - \gamma p+a p+a k\omega^{kl}\alpha o p+a] \); (12) \(-\frac{1}{3}[\alpha lkk\alpha o p+a p+a \alpha o p+a] \); (19) \( \omega^{vo}[\alpha l p+a p+a - \gamma p+a p+a k\omega^{kl}][\alpha o p+a p+b + 2\alpha o p+a p+b] \); (20) \( \omega^{vo}[\alpha l p+a p+a - \gamma p+a p+a k\omega^{kl}][\alpha o p+a p+b + 2\alpha o p+a p+b] \);

and the expectation of terms (17) and (21) are of \( O(n^{-3}) \).

Now, let \( J_{1}^{l{o}} \) be the sum of the expectations of the terms (1)-(3) and (8) multiplied by \( n^2 \), \( J_{2}^{l{o}} \) be that of (5), (7), (10) and (14), \( J_{3}^{l{o}} \) be that of (11), (12), (16) and (18), \( J_{4}^{l{o}} \) be that of (6) and (20), \( J_{5}^{l{o}} \) be that of (13) and (15) and finally \( J_{6}^{l{o}} \) be that of (9) and (19). Extensive
derivations show that, with a remainder terms of $O(n^{-1})$,

\[
J_1^{lo} = \frac{1}{4}(\alpha^{lk}k\alpha^{omm} + \alpha^{lk}m\alpha^{okm} + \alpha^{lo}k - \delta^{lo}) + \alpha^{lo} p+a p+a + \alpha^{l} p+a p+a \alpha^{o} p+b p+b
\]
\[
+ \alpha^{l} p+a p+b \alpha^{o} p+a p+b + \frac{1}{2}(\alpha^{lk}k\alpha^{omm} + 2\alpha^{lk}m\alpha^{okm}) - \frac{1}{3}(\alpha^{lk}k\alpha^{omm} + 2\alpha^{lk}m\alpha^{okm})
\]
\[
= \frac{1}{4}(\alpha^{lo}k - \delta^{lo}) + \frac{1}{36}\alpha^{lk}k\alpha^{omm} - \frac{7}{36}\alpha^{lk}m\alpha^{okm} + \alpha^{lo} p+a p+a + \alpha^{l} p+a p+a \alpha^{o} p+b p+b
\]
\[
+ \alpha^{l} p+a p+b \alpha^{o} p+a p+b
\]

\[
J_2^{lo} = [\alpha^{kl} p+a - \frac{1}{2} \gamma^p p+a m n\omega^{mk} \omega^{nl}] [\alpha^{ko} p+a - \frac{1}{2} \gamma^p p+a m n\omega^{mk} \omega^{nl}]
\]
\[
+ \frac{1}{2}(\alpha^{lk} k\alpha^{o} p+a p+a + \alpha^{lk} k\alpha^{o} p+a p+a)^{[2]} - \frac{1}{2}(\alpha^{lk} k\alpha^{o} p+a - \frac{1}{2} \gamma^p p+a m n\omega^{mk} \omega^{nl}] [\alpha^{ko} p+a]^{[2]}
\]
\[
- [\alpha^{kl} k\alpha^{o} p+a - \frac{1}{2} \gamma^p p+a m n\omega^{mk} \omega^{nl}] [\alpha^{ko} p+a]^{[2]}
\]
\[
= -\alpha^{lk} k\alpha^{o} p+a + \frac{1}{2} \gamma^p p+a m n\omega^{mk} \omega^{nl} \alpha^{ok} p+a)^{[2]}
\]
\[
+ \frac{1}{2} \gamma^p p+a m n\gamma^p p+a m n\omega^{mk} \omega^{nl} \alpha^{omm} \alpha^{kl}
\]

\[
J_3^{lo} = -\frac{1}{2}(\alpha^{l} p+a p+a - \gamma^p a p+a k, \omega^{kl}) [\alpha^{omm} \omega^{kl} p+a p+a)^{[2]}
\]
\[
+ \frac{1}{2} \alpha^{omm} \omega^{kl} p+a p+a, \omega^{kl} [\alpha^{l} p+a p+a - \gamma^p a p+a, \omega^{kl}]^{[2]}
\]
\[
= -\frac{1}{2} \alpha^{l} p+a p+a, \omega^{kl} [\alpha^{l} p+a p+a - \gamma^p a p+a, \omega^{kl}]^{[2]}
\]

\[
J_4^{lo} = \frac{1}{2} \gamma^p a p+a, \omega^{kl} [\alpha^{l} p+a p+a - \gamma^p a p+a, \omega^{kl}]
\]
\[
+ \frac{1}{2} \alpha^{l} p+a p+a, \omega^{kl} \gamma^p a p+a, \omega^{kl}
\]
\[
+ \gamma^p a p+a, \omega^{kl} [\alpha^{l} p+a p+a - \gamma^p a p+a, \omega^{kl}]
\]
\[
+ \omega^{kl} [\alpha^{l} p+a p+a - \gamma^p a p+a, \omega^{kl}] [\gamma^p a p+a, \omega^{kl}]
\]
\[
= \frac{1}{2} \gamma^p a p+a, \omega^{kl} [\alpha^{l} p+a p+a - \gamma^p a p+a, \omega^{kl}]
\]
\[
+ \gamma^p a p+a, \omega^{kl} [\gamma^p a p+a, \omega^{kl}]
\]
\[
= \frac{1}{2} \gamma^p a p+a, \omega^{kl} [\gamma^p a p+a, \omega^{kl}]
\]
\[
= \frac{1}{2} \gamma^p a p+a, \omega^{kl} [\gamma^p a p+a, \omega^{kl}]
\]
\[
J_5^{lo} = -\frac{1}{2} \gamma^p a p+a, \omega^{kl} [\alpha^{l} p+a p+a + \gamma^p a p+a, \omega^{kl}]
\]
\[
+ \frac{1}{2} \alpha^{l} p+a p+a, \omega^{kl} [\gamma^p a p+a, \omega^{kl}]
\]
\[
= -\frac{1}{2} \gamma^p a p+a, \omega^{kl} [\alpha^{l} p+a p+a + \gamma^p a p+a, \omega^{kl}]
\]
\[
+ \frac{1}{2} \alpha^{l} p+a p+a, \omega^{kl} [\gamma^p a p+a, \omega^{kl}]
\]
Let \( J \) be the leading order term of the expected value of all the 21 terms multiplied by \( n^2 \).

Then,

\[
J^{lo} = \frac{1}{4} (\alpha^{lkk} - \delta^{lo}) + \frac{1}{36} \alpha^{lkk} \alpha^{omm} - \frac{7}{36} \alpha^{km} \alpha^{okm} + \alpha^{lo} p+a p+a \\
- \alpha^{l} p+a p+b, o p+a p+b - \omega^{kl} \gamma^{p+a p+a, k} [2] + \gamma^{p+a p+b, k} \alpha^{o} p+a p+b, \omega^{kl} [2] \\
+ \frac{1}{2} \gamma^{p+a, a, m} \omega^{mk} \omega^{nl} \omega^{m'k} \omega^{n'l} - \frac{1}{2} \gamma^{p+a, mn} \gamma^{k p+a, v} \omega^{mk} \omega^{nl} \omega^{vo} [2] \\
(A.13)
\]

In summary, we have

\[
E(R_2^2 R_3^2) = n^{-2} J^{lo} + O(n^{-3}).
\]

Please note in the expressions for \( J^{lo} \) and \( J^{lo} \), there are terms like \( \omega^{kl} \). Although the superscript \( l \) is repeated, it does not imply summation over that superscript.
We also need to compute $E(R_2R_1^o) + E(R_3R_1^r)$. It may be shown that, with remainder terms of $O(n^{-1})$,

\[
\begin{align*}
E(R_2R_1^o) &= 5\alpha l k p + a + \frac{3}{5} \alpha l a k k p + a + \frac{2}{5} \alpha l a k k p + a + \alpha l p + a + \beta p + b \alpha o p + a + b + \\
E(R_3R_1^r) &= 5\alpha l o p + a + \alpha l p + a + \frac{1}{2} \alpha l o k p + a + \frac{1}{2} \alpha l a k k p + a + \frac{1}{2} \alpha l a k k p + a + b + \\
E(R_1R_3R_1^o) &= 5\alpha l k p + a + \frac{3}{5} \alpha l a k k p + a + \frac{2}{5} \alpha l a k k p + a + \frac{1}{2} \alpha l o p + a + \frac{1}{2} \alpha l o k p + a + \alpha l a k k p + a + b + \\
E(R_1R_3R_1^r) &= 5\alpha l o p + a + \alpha l p + a + \frac{1}{2} \alpha l o k p + a + \alpha l a k k p + a + b + \\
E(R_3R_4R_1^o) &= 5\alpha l k p + a + \frac{3}{5} \alpha l a k k p + a + \frac{2}{5} \alpha l a k k p + a + \frac{1}{2} \alpha l o p + a + \frac{1}{2} \alpha l o k p + a + \alpha l a k k p + a + b + \\
E(R_3R_4R_1^r) &= 5\alpha l o p + a + \alpha l p + a + \frac{1}{2} \alpha l o k p + a + \alpha l a k k p + a + b + \\
E(R_3R_5R_1^o) &= 5\alpha l k p + a + \frac{3}{5} \alpha l a k k p + a + \frac{2}{5} \alpha l a k k p + a + \frac{1}{2} \alpha l o p + a + \frac{1}{2} \alpha l o k p + a + \alpha l a k k p + a + b + \\
E(R_3R_5R_1^r) &= 5\alpha l o p + a + \alpha l p + a + \frac{1}{2} \alpha l o k p + a + \alpha l a k k p + a + b + \\
E(R_3R_6R_1^o) &= 5\alpha l k p + a + \frac{3}{5} \alpha l a k k p + a + \frac{2}{5} \alpha l a k k p + a + \frac{1}{2} \alpha l o p + a + \frac{1}{2} \alpha l o k p + a + \alpha l a k k p + a + b + \\
E(R_3R_6R_1^r) &= 5\alpha l o p + a + \alpha l p + a + \frac{1}{2} \alpha l o k p + a + \alpha l a k k p + a + b + \\
\end{align*}
\]
In light of (A.14) and (A.15), we have

\[ n^2 E(R_2^l R_1^o) = -\frac{1}{2} (\alpha_{lo}^{k;k} - \delta^{lo}) - \alpha^{lo} \ p + a \ p + a + \frac{1}{3} \alpha^{lkm} \alpha^{okm} + \alpha^{l} \ p + a \ p + b \ p + b + \alpha^{o} \ p + a \ p + b + \alpha^{l} \ p + a \ p + o \ p + a + \alpha^{ok} \ p + a \ p + a \ p + a + \frac{1}{2} \omega^{n;l} \omega^{m} \gamma_{p+a;oo;p+a;k} + \frac{1}{2} \omega^{n;l} \omega^{m} \gamma_{p+a;0+;a;k} \]

In summary,

(A.15) \[ E(R_2^l R_1^o) + E(R_3^l R_4^o) = n^{-2} K^{lo} + O(n^{-3}) \]

where

\[ K^{lo} = \frac{1}{8} (\alpha_{lok}^{k;k} + \delta^{lo}) - \frac{1}{8} \alpha_{lkm}^{m;k} - \frac{1}{8} \alpha_{lok}^{k;k} p + a \ p + a - \frac{1}{2} \omega^{m} \gamma_{p+a;oo;p+a;k} - \frac{1}{2} \omega^{n;l} \omega^{m} \gamma_{p+a;0+;a;k} \]

In light of (A.14) and (A.15), we have

(A.16) \[ \text{cum}(R^l, R^o) = n^{-1} \delta^{lo} + n^{-2} \Delta^{lo} + O(n^{-3}) \]

where

(A.17) \[ \Delta^{lo} = K^{lo}[2] + J^{lo} - \mu^{l} \mu^{o} \]

\[ = \frac{1}{2} \omega^{lok} + \alpha^{lo} \ p + a \ p + a - \frac{1}{3} \alpha^{lkm} \alpha^{okm} \]
which means the sum of the last three terms in (A.21) is negligible.

\[ \alpha_{18} \]

The joint fourth-order cumulants of 

\[ R_l, R_k, R_m, R_n \]

\[ \text{cum}(R_l, R_k, R_m, R_n) = E(R_l R_k R_m R_n) - E(R_l R_k)E(R_m R_n) \]
\[ - E(R_l^2)E(R_k R_m R_n) + 2E(R_l R_k)E(R_m R_n) \]
\[ - 6E(R_l R_k)E(R_m)E(R_n) \]

(A.21)

\[ = E(R_l^2 R_k R_m R_n) + E(R_l R_k R_m^2 R_n) + E(R_l R_k R_m R_n^2) + E(R_l R_k R_m R_n) \]
\[ + E(R_l R_k R_m^2 R_n) + E(R_l R_k R_m R_n^2) + E(R_l R_k R_m R_n) \]
\[ - E(R_l^2 R_k R_m R_n) - E(R_l R_k R_m^2 R_n) - E(R_l R_k R_m R_n^2) + O(n^{-4}) \]

From (A.11), we immediately have

\[ E(R_l^2) \{ E(R_l R_k R_m R_n) + E(R_l R_k R_m^2 R_n) + E(R_l R_k R_m R_n^2) + E(R_l R_k R_m R_n) \} = O(n^{-4}) \]

which means the sum of the last three terms in (A.21) is negligible.

To facilitate easy expressions, let us define

\[ t_1 = \alpha_{kmn}, t_2 = \delta_{lkm} \delta_{mn} + \delta_{lm} \delta_{kn} + \delta_{ln} \delta_{km} \]
\[ t_3 = \alpha^{lkn} \alpha^{mn} + \alpha^{lkn} \alpha^{mno} + \alpha^{lmn} \alpha^{koo} + \alpha^{kmn} \alpha^{l00}, \]
\[ t_4 = \alpha^{lko} \alpha^{mno} + \alpha^{lmo} \alpha^{kno} + \alpha^{lmo} \alpha^{k0n}, \]
\[ t_5 = \alpha^{kmn} \alpha^l p+a p+a + \alpha^{lmn} \alpha^k p+a p+a + \alpha^{lkn} \alpha^m p+a p+a + \alpha^{lkm} \alpha^n p+a p+a, \]
\[ t_6 = \alpha^{lk} p+a \alpha^{mn} p+a + \alpha^{lmo} p+a \alpha^{kn} p+a + \alpha^{ln} p+a \alpha^{km} p+a. \]

It is relatively easy to show that

\[(A.22) \quad E(R_1^4 R_1^4 R_1^4 R_1^4) - E(R_1^4 R_1^4) E(R_1^4 R_1^4)[3] = n^{-3}(t_1 - t_2) + O(n^{-4}).\]

To derive \( E(R_1^4 R_1^4 R_1^4 R_1^4)[4] - E(R_1^4 R_1^4) E(R_1^4 R_1^4)[12], \) we notice that

\[-\frac{1}{2} \{E(A^{l0} A^o A^k A^m A^n)[4] - E(A^{l0} A^o A^k) E[A^m A^n][12] \} = n^{-3} I_2 + O(n^{-4}),\]

where \( I_2 = -6 t_1 + 2 t_2 - \frac{1}{2} t_3 - 2 t_4. \)

\[-\{E(A^{l p+a} A^{p+a} A^k A^m A^n)[4] - E(A^{l p+a} A^{p+a} A^k) E[A^m A^n][12] \} = n^{-3} I_3 + O(n^{-4}),\]

where \( I_3 = -(t_5 + 4 t_6), \)

\[-\frac{1}{3} \alpha^{l0} \{E(A^o A^v A^k A^m A^n)[4] - E(A^o A^v A^k) E(A^m A^n)[12] \} = n^{-3} I_4 + O(n^{-4}),\]

where \( I_4 = \frac{4}{3} t_3 + \frac{8}{3} t_4. \)

\[ \omega^{ol} \{E(C^{p+a,0} A^{p+a} A^k A^m A^n)[4] - E(C^{p+a,0} A^{p+a} A^k) E(A^m A^n)[12] \} = n^{-3} I_5 + O(n^{-4}), \]

where \( I_5 = \omega^{ol} (\gamma^{p+a,0} \alpha^{kn} + \gamma^{p+a,0} \alpha^{mn} p+a + \gamma^{p+a,0} \alpha^{kn} p+a + \gamma^{p+a,0} \alpha^{kn} p+a)[4], \)

\[ [\alpha^{l0} p+a - \frac{1}{2} \gamma^{p+a,0} p+a \omega^{ol} \omega^{ol}] \{E(A^{p+a} A^o A^k A^m A^n)[4] - E(A^{p+a} A^o A^k) E(A^m A^n)[12] \} = n^{-3} I_6 + O(n^{-4}), \]

where \( I_6 = 4 t_6 - [\omega^{m'l} \omega^{m'l} \gamma^{p+a,0} p+a][6] \) and

\[ \{[\alpha^{l0} p+a p+b - \gamma^{p+a,0} p+a \omega^{ol} \{E[A^{p+a} A^{p+b} A^k A^m A^n][4] - E[A^{p+a} A^{p+b} A^k] E[A^m A^n][12] \} = n^{-3} I_7 + O(n^{-4}), \]

where \( I_7 = t_5 - \gamma^{p+a,0} p+a[\omega^{ol} \alpha^{kn}][4]. \)
In summary, we have

\[
E(R_k^l R_k^m R_k^n) = E(R_k^l R_k^m)E(R_k^n) = n^{-3} \{ \omega_{ol}(\gamma_{p+a,om}\alpha_{mn} p+a + \gamma_{p+a,om}\alpha_{kn} p+a + \gamma_{p+a,om}\alpha_{km} p+a) \} \{ 6 \} - 6t_1 + 2t_2 - \frac{1}{6}t_3 + \frac{2}{3}t_4 - [\omega_{ol}(\gamma_{p+a,m'n'}\alpha_{mn} p+a) \} \{ 6 \} \} + O(n^{-4}).
\]

(A.23)

Next we need to calculate \( E(R_k^l R_k^m R_k^n) \) for \( j = 1, \cdots, 21 \). The expressions for \( I_{j+7} \) are

\[
I_8 = 3t_1 - t_2 + \frac{3}{2}t_3 + 3t_4, \quad I_9 = 4t_6, \quad I_{10} = \frac{2}{3}t_3 + \frac{16}{3}t_4,
\]

\[
I_{11} = [\omega_{ol}(\gamma_{p+a,m'n'}\omega_{m'n'}\omega_{mn})] \{ 6 \},
\]

\[
I_{12} = [(\alpha_{kn} p+a - \frac{1}{2}\gamma_{p+a,om}\omega_{mn}(\omega_{m'n'}\omega_{mn})\omega_{mn})\omega_{mn} \} \{ 6 \},
\]

\[
I_{13} = 0, \quad I_{14} = 3t_5 + 4t_6, \quad I_{15} = -2t_3 - \frac{16}{3}t_4,
\]

\[
I_{16} = \{ \omega_{ol}(\gamma_{p+a,om}\alpha_{mn} p+a + \omega_{ok}\alpha_{mn}) \} \{ 6 \},
\]

\[
I_{17} = -4t_6 + \frac{1}{2}(\gamma_{p+a,m'n'}\omega_{m'n'}(\omega_{m'n'}\omega_{mn} \} \{ 6 \},
\]

\[
I_{18} = [\omega_{ol}(\gamma_{p+a,om}\alpha_{mn} p+a + \omega_{ok}\alpha_{mn}) \} \{ 6 \}, \quad I_{19} = 2t_5,
\]

\[
I_{20} = [\omega_{ol}(\gamma_{p+a,om}\alpha_{mn} p+a + \omega_{ok}\alpha_{mn}) \} \{ 6 \},
\]

\[
I_{21} = -8t_6 + \frac{1}{2}(\gamma_{p+a,m'n'}\omega_{m'n'}(\omega_{m'n'}\omega_{mn} \} \{ 6 \},
\]

32
\[ E(R_2 R_2^m R_1^m R_1)[6] - E(R_2 R_2^k R_1^m R_1)[6] = n^{-3} \sum_{j=1}^{21} I_{j+7} + O(n^{-4}) \]

\[ = n^{-3} \{ 3t_1 - t_2 + \frac{1}{6} t_3 - \frac{2}{5} t_4 + [\omega^{m'} \omega^{n'} + \gamma^{p+a, o, m'} \omega^{m n'} \alpha^{k n p+a}] [6] \]

\[ - \frac{1}{2} [\omega^{k'} \omega^{n'} + \gamma^{p+a, o, k'} \omega^{m n'} \alpha^{l m p+a}] [6] \]

\[ + \omega^{m' n'} \alpha^{l n p+a} + \gamma^{p+a, o, m} \omega^{l m n'} [6] \]

\[ + [\omega^{m'} \omega^{n'} \alpha^{k n p+a} + \gamma^{p+a, o, m} \omega^{m n'} \alpha^{l m p+a}] [6] \]

\[ = 1 \frac{1}{2} [\omega^{l' m' n' k'} (\omega^{m' n' k'} + \omega^{n' m' k'}) + \gamma^{p+a, l' m' n' k'} + \gamma^{p+a, m' n' k'}] [6] + O(n^{-4}). \]  

(A.24)

Finally, we need to derive \( E(R_3 R_1 R_1^m R_1)[4] - E(R_3 R_1^k R_1^m R_1)[12] \). Due to the fact that \( E(A^{p+a} A^l) = 0 \) for \( l \in \{1, \cdots, p\} \), there is no need to compute the terms in \( R_3 \) involving \( A^{p+a} \). In particularly,

\[ E(R_3 R_1 R_1^m R_1)[4] - E(R_3 R_1^k R_1^m R_1)[12] = n^{-3} I_{29} + O(n^{-4}) \]

where \( I_{29} = 2t_1 - \frac{1}{6} t_4 \),

\[ E(A^{l' p+a} A^{p+a} A^{p+a} R_1 R_1^m R_1)[4] - E(A^{l' p+a} A^{p+a} A^{p+a} R_1 R_1^m)[12] = n^{-3} I_{30} + O(n^{-4}) \]

where \( I_{30} = 4t_6 \),

33
\[-\frac{1}{2} \omega^{n' l'} \omega^{k' m'} \{ E(C^{p+a,n'} C^{p+a,k'} A^{n'} R_k^1 R_1^m R_n^m) [4] - E(C^{p+a,n'} C^{p+a,k'} A^{n'} R_k^1) E(R_1^m R_1^m) [12] \} = n^{-3} I_{31} + O(n^{-4})\]

where \( I_{31} = \{ \omega^{n' l'} \omega^{k' m'} [\gamma^{p+a,k';n} p+a,n';m + \gamma^{p+a,n';n} p+a,k';m] \} \) [6],

\[-\frac{1}{2} \omega^{k' n'} \omega^{l' o'} \omega^{m' v'} \{ E(C^{p+a,m'} A^{n'} A^{n'} A^{n'} R_k^1 R_1^m R_n^m) [4] - E(C^{p+a,m'} A^{n'} A^{n'} R_k^1) E(R_1^m R_1^m) [12] \} = n^{-3} I_{32} + O(n^{-4}),\]

where \( I_{32} = \frac{1}{2} \{ \gamma^{p+a,m'n'} (\omega^{k' n'} + \omega^{l' o'}) (\omega^{m' v'} A^{n'} A^{n'} A^{n'} A^{n'}) [4] - E(C^{p+a,m'} A^{n'} A^{n'} A^{n'} A^{n'}) E(R_1^m R_1^m) [12] \} = n^{-3} I_{34} + O(n^{-4}),\]

where \( I_{34} = -8t_6, \)

\[-E(A^{n'} A^{n'} A^{n'} A^{n'}) E(R_1^m R_1^m) [12] \} = n^{-3} I_{34} + O(n^{-4}),\]

where \( I_{34} = 4t_6 - \frac{1}{2} \{ \omega^{l' o'} \omega^{m' m'} (\omega^{n' l'} \omega^{k' n'} + \omega^{n' o'} \omega^{k' m'} \gamma^{p+a,l'} m' \gamma^{p+a,n'} m') [6].\]

In summary, we have

\[E(R_1^m R_1^m R_1^m) [4] - E(R_1^m R_1^m) E(R_1^m R_1^m) [12] = n^{-3} \sum_{j=1}^{6} I_{j+28} + O(n^{-4})\]

\[-E(R_1^m R_1^m R_1^m R_1^m) [4] - E(R_1^m R_1^m R_1^m R_1^m) [12] = n^{-3} \sum_{j=1}^{6} I_{j+28} + O(n^{-4})\]

Combining (A.22), (A.23), (A.24) and (A.25) it may be shown that

(A.26) \( \text{cum}(R_k^l, R_k^m, R_k^m, R_k^m) = O(n^{-4}). \)

Edgeworth Expansion for \( r(\theta_0) \). We first derive an Edgeworth expansion for the distribution of \( n^{1/2} R \). Let \( \kappa_j \) be the \( j \)-th order joint cumulant of \( n^{1/2} R \). From (A.10), (A.16), (A.12) and (A.26),

\[ \kappa_1 = n^{-1/2} \mu + O\left(n^{-3/2}\right), \quad \kappa_2 = I_p + n^{-1} \Delta + O(n^{-3}), \]

where

\[ p = \frac{1}{2}, \quad q = \frac{1}{2}, \quad r = \frac{1}{2}, \quad s = \frac{1}{2}, \quad t = \frac{1}{2}, \quad u = \frac{1}{2}, \quad v = \frac{1}{2}, \quad w = \frac{1}{2}, \quad x = \frac{1}{2}, \quad y = \frac{1}{2}, \quad z = \frac{1}{2}. \]
\[ \kappa_3 = O(n^{-3/2}), \quad \kappa_4 = O(n^{-2}), \]

where \( I_p \) is the \( p \times p \) identity matrix, \( \mu = (\mu^1, \cdots, \mu^p)^T \) with \( \mu^k = -\frac{1}{6} \alpha^{kk} \) and \( \Delta = (\Delta^{kk})_{p \times p} \).

Let

\[ \bar{U}_A = \begin{pmatrix} A^1, \cdots, A^r, A^{11}, \cdots, A^{rr}, A^{111}, \cdots, A^{rrr} \end{pmatrix}^T, \]
\[ \bar{U}_C = \begin{pmatrix} C^{11}, \cdots, C^{1p}, \cdots, C^{r1}, \cdots, C^{rp}, C^{111}, \cdots, C^{rrr} \end{pmatrix}^T \]

and \( \bar{U} = (\bar{U}_A^T, \bar{U}_C^T)^T \) is a vector of means. From (2.10), (2.11) and the expansion for \( R_3 \) given in Appendix A.3, the signed square root \( n^{1/2} R \) can be expressed as a smooth function of \( U \), namely there exists a smooth function \( h \) such that \( n^{1/2} R = h(\bar{U}) \). We can then use the results given in Bhattacharya and Ghosh (1978) to formally establish Edgeworth expansion for the distribution of \( n^{1/2} R \) under conditions (2.1). In particular, let \( \mathcal{B} \) be a class of Borel sets in \( \mathbb{R}^p \) satisfying

\[ \sup_{B \in \mathcal{B}} \int_{\partial B} \phi(v) dv = O(\epsilon), \quad \epsilon \downarrow 0, \]

where \( \partial B \) and \( (\partial B)^\epsilon \) are the boundary of \( B \) and \( \epsilon \)-neighborhood of \( \partial B \) respectively. A formal Edgeworth expansion for the distribution function of \( n^{1/2} R \) is

\[ \sup_{B \in \mathcal{B}} |P(n^{1/2} R \in B) - \int_B \pi(v) \phi(v) dv| = O(n^{-3/2}), \]

where \( \pi(v) = 1 + n^{-1/2} \mu^T v + \frac{1}{2} n^{-1} \{ v^T (\mu \mu^T + \Delta) v - \text{tr} (\mu \mu^T + \Delta) \} \), \( \phi(v) \) is the \( p \)-dimensional standard normal density, and \( \text{tr}() \) is the trace operation for square matrices.

Let \( H = (h_{ij})_{p \times p} =: \mu \mu^T + \Delta \). By the symmetry of \( \phi(v) \) we have

\[
\begin{align*}
P\{ \ell(\beta) < c_\alpha \} & = P\{ (n^{1/2} R)^T (n^{1/2} R) < c_\alpha \} + O(n^{-3/2}) \\
& = \int_{\|v\| < c_\alpha^{1/2}} \pi(v) \phi(v) dv + O(n^{-3/2}) \\
& = P(\chi_p^2 < c_\alpha) + \frac{1}{2} n^{-1} \int_{\|v\| < c_\alpha^{1/2}} \{ \sum_{i=1}^p h_{ii} v_i^2 - 1 + \sum_{i \neq j} h_{ij} v_i v_j \} \phi(v) dv \\
& \quad + O(n^{-3/2}) \\
& = \alpha - p^{-1} B_e c_\alpha f_{p}(c_\alpha) n^{-1} + O(n^{-3/2})
\end{align*}
\]

(A.27)
where $B_c = \sum_{i=1}^{p} h_{ii} = \sum_{i=1}^{p} (\mu_i^j + \Delta^i_i)$. Due to the fact that an even order Hermite polynomial is an even function, the error term in (A.27) is in fact $O(n^{-2})$. This completes the proof.

**Proof of Theorem 2.** Based on the Edgeworth expansion given in Theorem 1 and follow standard derivations, for instance those given in Chen (1993), we can readily establish the theorem.

**References**


metrics, 117, 55-93.


Department of Statistics, Iowa State University, Ames, Iowa 50011-1210, USA
Email: songchen@iastate.edu

Department of Mathematics, Beijing Normal University, 100875, China
Email: hjcui@bnu.edu.cn
Table 1. Empirical coverage (in percentage) and averaged length of the EL confidence interval $I_\alpha$, the Bartlett corrected (BC) EL interval $I_{\alpha, bc}$ and the direct Bootstrap (BT) calibrated confidence interval $I_{\alpha, bt}$ for the normal mean example.

(a) $N(0, 1)$

<table>
<thead>
<tr>
<th>Nominal level Sample Size</th>
<th>90% Coverage</th>
<th>90% Length</th>
<th>95% Coverage</th>
<th>95% Length</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>EL</td>
<td>BC</td>
<td>BT</td>
<td>EL</td>
</tr>
<tr>
<td>20 coverage</td>
<td>84.66</td>
<td>89.40</td>
<td>86.88</td>
<td>91.12</td>
</tr>
<tr>
<td>length</td>
<td>0.636</td>
<td>0.817</td>
<td>0.791</td>
<td>0.757</td>
</tr>
<tr>
<td>30 coverage</td>
<td>86.30</td>
<td>89.60</td>
<td>87.80</td>
<td>92.60</td>
</tr>
<tr>
<td>length</td>
<td>0.552</td>
<td>0.662</td>
<td>0.640</td>
<td>0.659</td>
</tr>
<tr>
<td>40 coverage</td>
<td>86.47</td>
<td>89.48</td>
<td>87.98</td>
<td>93.49</td>
</tr>
<tr>
<td>length</td>
<td>0.492</td>
<td>0.558</td>
<td>0.548</td>
<td>0.588</td>
</tr>
<tr>
<td>60 coverage</td>
<td>86.70</td>
<td>89.00</td>
<td>87.90</td>
<td>93.50</td>
</tr>
<tr>
<td>length</td>
<td>0.448</td>
<td>0.492</td>
<td>0.437</td>
<td>0.536</td>
</tr>
</tbody>
</table>

(b) $N(1, 2)$

<table>
<thead>
<tr>
<th>Nominal level Sample Size</th>
<th>90% Coverage</th>
<th>90% Length</th>
<th>95% Coverage</th>
<th>95% Length</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>EL</td>
<td>BC</td>
<td>BT</td>
<td>EL</td>
</tr>
<tr>
<td>20 coverage</td>
<td>80.02</td>
<td>87.99</td>
<td>85.57</td>
<td>86.78</td>
</tr>
<tr>
<td>length</td>
<td>0.656</td>
<td>0.898</td>
<td>0.854</td>
<td>0.781</td>
</tr>
<tr>
<td>30 coverage</td>
<td>82.85</td>
<td>89.47</td>
<td>88.16</td>
<td>88.87</td>
</tr>
<tr>
<td>length</td>
<td>0.6547</td>
<td>0.7740</td>
<td>0.752</td>
<td>0.739</td>
</tr>
<tr>
<td>40 coverage</td>
<td>85.77</td>
<td>89.94</td>
<td>88.92</td>
<td>91.77</td>
</tr>
<tr>
<td>length</td>
<td>0.496</td>
<td>0.590</td>
<td>0.572</td>
<td>0.592</td>
</tr>
<tr>
<td>60 coverage</td>
<td>87.47</td>
<td>90.28</td>
<td>89.68</td>
<td>92.69</td>
</tr>
<tr>
<td>length</td>
<td>0.411</td>
<td>0.454</td>
<td>0.446</td>
<td>0.492</td>
</tr>
</tbody>
</table>
Table 2. Empirical coverage (in percentage) and averaged length of the EL confidence interval $I_{\alpha}$ and the Bartlett corrected (BC) EL interval $I_{\alpha, \text{bc}}$ for the panel data model (4.1)

(a) $\theta = 0.5$

<table>
<thead>
<tr>
<th>Nominal level</th>
<th>90%</th>
<th>95%</th>
</tr>
</thead>
<tbody>
<tr>
<td>Sample Size</td>
<td>EL</td>
<td>BC</td>
</tr>
<tr>
<td>50 coverage</td>
<td>81.7</td>
<td>87.3</td>
</tr>
<tr>
<td>length</td>
<td>0.603</td>
<td>0.723</td>
</tr>
<tr>
<td>100 coverage</td>
<td>87.4</td>
<td>89.9</td>
</tr>
<tr>
<td>length</td>
<td>0.425</td>
<td>0.461</td>
</tr>
</tbody>
</table>

(b) $\theta = 0.9$

<table>
<thead>
<tr>
<th>Nominal level</th>
<th>90%</th>
<th>95%</th>
</tr>
</thead>
<tbody>
<tr>
<td>Sample Size</td>
<td>EL</td>
<td>BC</td>
</tr>
<tr>
<td>50 coverage</td>
<td>84.4</td>
<td>89.6</td>
</tr>
<tr>
<td>length</td>
<td>0.545</td>
<td>0.648</td>
</tr>
<tr>
<td>100 coverage</td>
<td>87.3</td>
<td>89.1</td>
</tr>
<tr>
<td>length</td>
<td>0.373</td>
<td>0.401</td>
</tr>
</tbody>
</table>